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Multisensory integration operates on correlated input from unimodal transients channels

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Abstract

Audiovisual information reaches the brain via both sustained and transient input channels, representing signals' intensity over time or changes thereof, respectively. To date, it is unclear to what extent transient and sustained input channels contribute to the combined percept obtained through multisensory integration. Based on the results of two novel psychophysical experiments, here we demonstrate the importance of the transient (instead of the sustained) channel for the integration of audiovisual signals. To account for the present results, we developed a biologically-inspired, general-purpose model for multisensory integration, the Multisensory Correlation Detectors, which combines correlated input from unimodal transient channels. Besides accounting for the results of our psychophysical experiments, this model could quantitatively replicate several recent findings in multisensory research, as tested against a large collection of published datasets. In particular, the model could simultaneously account for the perceived timing of audiovisual events, multisensory facilitation in detection tasks, causality judgments, and optimal integration. All-in-all, this study demonstrates that several phenomena in multisensory research that were previously considered unrelated, all stem from the integration of correlated input from unimodal transient channels.

eLife assessment

This **important** study evaluates a model for multisensory correlation detection, focusing on the detection of correlated transients in visual and auditory stimuli. Overall, the experimental design is sound and the evidence is **compelling**. The synergy between the experimental and theoretical aspects of the paper is strong, and the work will be of interest to both neuroscientists and psychologists working in the domain of sensory processing and perception.

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Introduction

Audiovisual stimuli naturally unfold over time, and their structures alternate intervals during which the signals remain relatively constant (such as the steady luminance of a light bulb) to sudden moments of change (when the bulb lights up). To efficiently process the temporal structure of incoming signals, the sensory systems of mammals (and other animal classes) rely on separate channels, encoding stimulus intensity through either sustained or transient responses (Benucci, Frazor, and Carandini 2007; Kim et al. 2011; Breitmeyer and Ganz 1976; Qin et al. 2007; Recanzone 2000; Ikeda and Wright 1972). These channels are known to originate early in the processing hierarchy (Recanzone 2000; Ikeda and Wright 1972), with sustained ones responding with constant neural firing to static input intensity, whereas transient channels respond with increased firing to any variations in input intensity (i.e., both increments and decrements, **Figure 1A**). On a functional level, the response of sustained and transient channels represents distinct dynamic stimulus information. Specifically, sustained responses represent the intensity of the stimulus, while transient responses represent changes in stimulus intensity. In terms of frequency response, sustained channels can be characterised as low-pass temporal filters (**Equation 10**), highlighting the low-frequency signal components. Transient channels, on the other hand, have a higher spectral tuning and can be characterised as band-pass temporal filters (**Equation 1**), signaling events or moments of stimulus change (Stigliani, Jeska, and Grill-Spector 2017).

Changing information over time is also critical for multisensory perception (Barry E. Stein 2012): when two signals from different modalities are caused by the same underlying event, they usually covary over time (like firecrackers' pops and blazes). A growing body of literature has now investigated human sensitivity and adaptation to temporal lags across the senses (Vroomen and Keetels 2010), and it is well established that both multisensory illusions (Sekuler, Sekuler, and Lau 1997; van Wassenhove, Grant, and Poeppel 2007; Samad et al. 2018) and Bayesian-optimal cue integration (e.g., Parise, Spence, and Ernst 2012) critically depend on synchrony and temporal correlation across the senses. However, multisensory integration does not operate on the raw sensory signals; these are systematically transformed during transduction and early neural processing. Therefore to understand multisensory integration, it is critical to figure out how unisensory signals are processed before feeding into the integration stage. Surprisingly, this fundamental question has received little attention in multisensory research.

Current evidence, however, suggests a prominence of transient channels in the percept resulting from multisensory integration. For example, Andersen and Mamassian (Andersen and Mamassian 2008) found that task-irrelevant increments or decrements in sound intensity equally facilitated the detection of both increments and decrements in the lightness of a visual display. Critically, such an effect only occurred when changes in the two modalities occurred in approximate temporal synchrony. Based on this independence of polarity of this crossmodal facilitation, where intensity increments and decrements produced similar perceptual benefits, the authors concluded that audiovisual integration relies primarily on unsigned transient stimulus information. The role of transient channels in audiovisual perception is further supported by fMRI evidence: Werner and Noppeney (Werner and Noppeney 2011) found that audiovisual interactions in the human brain only occurred during stimulus transitions, and demonstrated that transient onset and offset responses could be dissociated both anatomically and functionally (see also, Herdener et al. 2009). While these studies demonstrate the dominance of transient over sustained temporal channels in e.g. detection tasks as studied by Andersen and Mamassian (Andersen and Mamassian 2008), or the pattern of neural responses during passive observation of audiovisual stimuli as in Herdener et al (Herdener et al. 2009), to date, it is still unknown to what extent transient and

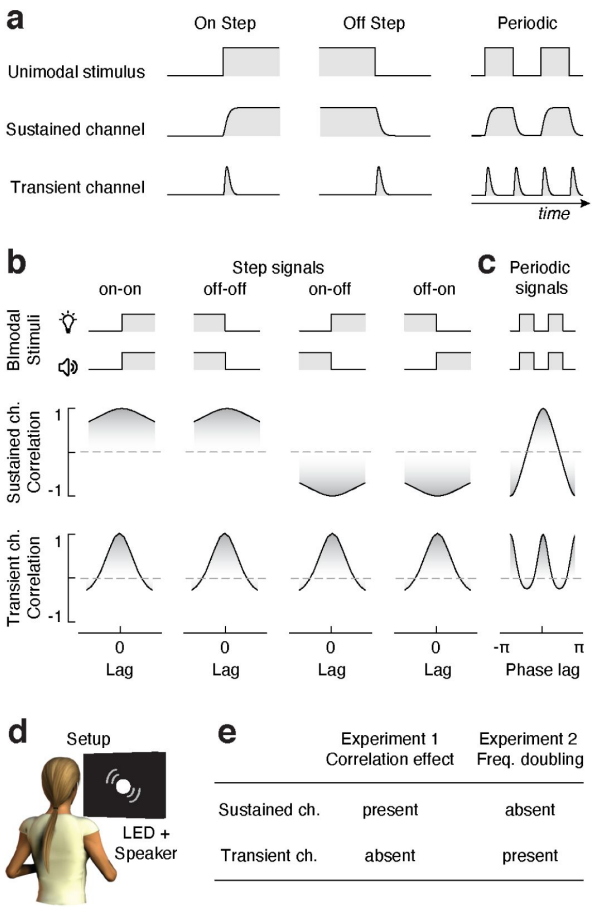


Figure 1.

Sustained vs transient channels.

A. Responses of sustained and transient channels to onset and offset step stimuli, and periodic signals comprising sequences of onsets and offsets. Note that while the sustained channels closely follow the intensity profile of the input stimuli, transient channels only respond to changes in stimulus intensity, and such a response is always positive, irrespective of whether stimulus intensity increases or decreases. Therefore, when presented with periodic signals, while the sustained channels respond at the same frequency as the input stimulus (frequency following), transient channels respond at a frequency that is twice that of the input (frequency doubling). **B.** Synchrony as measured from cross-correlation between pairs of step stimuli, as seen through sustained (top) and transient (bottom) channels (transient and sustained channels are simulated using Equations 1 and 10, respectively). Note how synchrony (i.e, correlation) for sustained channels peaks at zero lag when the intensity of the input stimuli changes in the same direction, whereas it is minimal at zero lag when the steps have opposite polarities (negatively correlated stimuli). Conversely, being insensitive to the polarity of intensity changes, synchrony for transient channels always peaks at zero lag. **C.** Synchrony (i.e., cross-correlation) of periodic onsets and offset stimuli as seen from sustained and transient channels. While synchrony peaks once (at zero phase shift) for sustained channels, it peaks twice for transient channels (at zero and pi radians phase shift), as a consequence of its frequency doubling response characteristic. **D.** Experimental apparatus: participants sat in front of a black cardboard panel with a circular opening, through which audiovisual stimuli were delivered by a white LED and a loudspeaker. **E.** Predicted effects of Experiments 1 and 2 depending on whether audiovisual integration relies on transient or sustained input channels. The presence of the effects of interest in both experiments or the lack thereof indicates an inconclusive result, not interpretable in the light of our hypotheses.

sustained channels affect the perceived timing of audiovisual events – such as the subjective synchrony of visual and auditory signals, which is arguably the primary determinant of multisensory integration.

To understand the effect of transient versus sustained channels in multisensory perception, we must first focus on the difference between their unimodal responses. For that, we can consider stimuli consisting of steps in stimulus intensity (e.g. (Andersen and Mamassian 2008). A schematic representation is shown in **Figure 1A**: while onset and offset step stimuli trigger identical unsigned transient responses, sustained responses differ across conditions. That is, given that sustained channels represent the magnitude of the stimulus, responses to onset and offset stimuli are negatively correlated. In signal processing, Pearson correlation is commonly used to assess the synchrony of two related signals, with a higher correlation representing higher synchrony (and similarity, (Wei 2006). Therefore, we can hypothesise that if multisensory time perception relies on sustained input channels, positively correlated audiovisual stimuli (e.g., onset or offset stimuli in both modalities) should be perceived as more synchronous than negatively correlated stimuli (i.e., onset in one modality, offset in the other). Conversely, if audiovisual synchrony relies on unsigned transients, positively and negatively correlated stimuli should appear equally synchronous (**Figure 1B**).

This hypothesis will be tested in Experiment 1, where systematic differences in perceived synchrony based on whether audiovisual signals are positively or negatively correlated, would provide evidence for a dominant role of sustained input channels in audiovisual temporal processing. A lack of systematic differences driven by stimulus correlation, however, would not necessarily imply a transient nature of audiovisual temporal processing: this would require additional evidence from an inverse experiment, one in which the transient hypothesis predicts an effect that the sustained hypothesis does not (**Figure 1E**). For that, we can consider audiovisual stimuli consisting of periodic onsets and offsets (e.g., square-wave amplitude modulation, see **Figure 1C**). While sustained channels respond at the same rate as the input, transient responses have a rate double that of the input signals. This phenomenon, known as frequency doubling, is commonly considered a hallmark of the contribution of transient channels in sensory neuroscience (e.g., Kim et al. 2011).

Frequency doubling, therefore, is a handle to assess the contribution of transient and sustained channels in audiovisual perception. Consider audiovisual stimuli defined by square-wave amplitude modulations with a parametric manipulation of crossmodal phase shift (**Figure 1C**). If audiovisual temporal integration relies on the correlation between sustained input channels, we can predict perceived audiovisual synchrony (i.e., correlation, Wei 2006) to peak just once: at zero phase shifts. Conversely, if multisensory temporal integration relies on transient channels, perceived audiovisual synchrony should peak twice: at zero and 180deg. phase lag (**Figure 1C**). Such a frequency doubling phenomenon can be easily assessed psychophysically, by measuring reported simultaneity as a function of audiovisual lags. Therefore, in a second psychophysical study, we rely on frequency doubling in audiovisual synchrony perception to assess whether multisensory integration relies on transient or sustained input channels.

Experiment 1: Step stimuli

To probe the effect of the correlation between unimodal step stimuli on the perceived timing of audiovisual events, eight participants (age range 22–35 years, four females) observed audiovisual signals consisting of lightness and acoustic intensity increments (on-step) and decrements (off-step). On-steps and off-steps were paired in all possible combinations, giving rise to four experimental conditions (both modalities on, both off, vision on with audio off, and vision off with audio on, see **Figure 1B**). The lag between visual and acoustic step events was parametrically manipulated using the method of constant stimuli (15 steps, between -0.4 and 0.4s). After the

stimulus presentation, participants performed a temporal order judgment (TOJ, which event came first, sound or light?) or a simultaneity judgment (SJ, were the stimuli synchronous or not?). TOJ and SJ tasks were run in different sessions occurring on different days. Although our original hypotheses (see, **Figure 1B-E**) do not make specific predictions for the TOJ task, for completeness we run such an experiment anyway, as its inclusion provides a more stringent test for our model (see "Modelling" section).

The audiovisual display used to deliver the stimuli consisted of a white LED in an on- (high-luminance) and off-state (low-luminance), and acoustic white noise, also in either on-(quiet) or off-state (loud). Sounds came from a speaker located behind the LED, and both the speaker and the LED were controlled using an audio interface to minimize system delay (**Figure 1D**, see (Parise and Ernst 2016). A white, sound transparent cloth was placed in front of the LED so that when the light was on, participants saw a white disk of 13° in diameter. Overall, each participant provided 600 responses in the TOJ task (15 lags, 10 repetitions, and 4 conditions) and additional 600 responses in the SJ task (again, 15 lags, 10 repetitions, and 4 conditions). The experiment was run in a dark, sound-attenuated booth, and the position of participants' heads was controlled using a chin- and a headrest. Participants were paid 8 Euro/hour. The experimental procedure was approved by the Ethics committee of the University of Bielefeld and was conducted in accordance with the declaration of Helsinki.

Results

To assess whether different combinations of on-step and off-step stimuli (and correlation thereof) elicit measurable psychophysical effects, we estimated both the point and the window of subjective simultaneity (PSS and WSS, that is, the delay at which audiovisual stimuli appear simultaneous, and the width of the window of simultaneity). For that, we fitted psychometric curves to both TOJ and SJ data, independently for each condition. Specifically, following standard procedures (Parise and Spence 2009), TOJs were statistically modelled as cumulative Gaussians, with 4 free parameters (intercept, slope, and two asymptotes, **Figure 2A**). The PSS was calculated as the lag at which TOJs were at chance level, whereas the window of simultaneity (WSS) was calculated as the half-difference between the lags eliciting 0.75 and 0.25 probability of audio-first responses. The SJs data, instead, were modeled as the difference of two cumulative Gaussians (Yarrow et al. 2011), leading to asymmetric bell-shaped psychometric functions (**Figure 2B**). The PSS was calculated as the lag at which perceived simultaneity was maximal, whereas the window of simultaneity was calculated as the half-width at half-maximum. Results are summarized in **Figure 2C-D** (see **Supplementary Figure S1** for individual data, and Supplementary Information for a correlation analysis of the PSS and WSS measured from TOJs vs SJs).

To assess whether the four experimental conditions differ in the point and window of perceived simultaneity (PSS & WSS), we run both non-parametric Friedman tests, with Bonferroni correction for multiple testing, and Bayesian repeated measures ANOVA Neither PSS nor WSS statistically differed across conditions, and this was true for both the TOJ and SJ data. Results are summarized in Table S2.

Discussion

The lack of a difference across conditions found in Experiment 1 implies that the on-step and off-step stimuli induced similar perceptual responses. Based on our original hypothesis, the present results argue against the dominance of sustained input channels in the perception of audiovisual events, as the synchrony (i.e., Pearson correlation, Wei 2006) of sustained signals, would otherwise have been affected by the crossmodal combination of on-step and off-step stimuli. As previously mentioned, while a lack of difference across conditions in Experiment 1 is necessary to infer a dominance of transient channels in audiovisual time perception, this evidence is not sufficient on its own. For that, we would need additional evidence in the form of the presence of

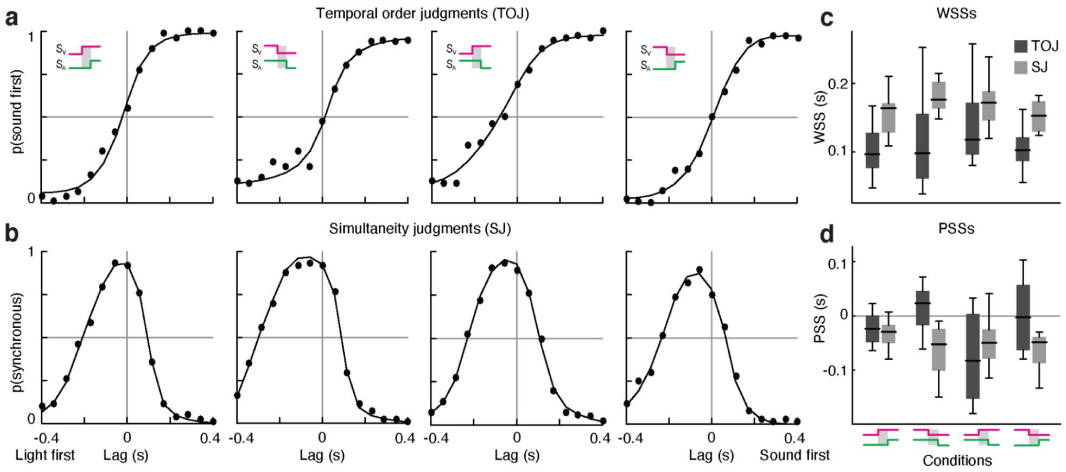


Figure 2.

Experiment 1, results.

A. Responses in the TOJ task and psychometric fits (averaged across participants) for the four experimental conditions. **B.** Responses in the SJ task and psychometric fits (averaged across participants) for the four experimental conditions. Each dot in panels A and B corresponds to 80 trials. **C.** Window of subjective simultaneity for each condition and task. **D.** Point of subjective simultaneity for each condition and task.

systematic effects, which are predicted from the operating principles of the transient channels, but not for sustained ones. Experiment 2 was designed to fulfill such a requirement, as it predicts a frequency-doubling effect in perceived synchrony for transient but not for sustained input channels.

Experiment 2: Periodic stimuli

Participants observed a periodic audiovisual stimulus consisting of a square-waved intensity envelope and performed a force-choice simultaneity judgment task. The carrier visual and auditory stimuli consisted of pink noise, delivered by a speaker and an LED (**Figure 1D**), which were switched on and off periodically in a square-wave fashion with a period of 2 s for a total duration of 6 s (so that three full audiovisual cycles were presented on each trial, **Figure 3A**). To prevent participants from just focusing on the start and endpoint of the signals, during the intertrial interval, the visual and auditory stimuli were set to a pedestal intensity level, so that during the trial, the square wave modulation was gradually ramped on and off, following a raised cosine profile with a duration of 6 s (**Figure 3A**).

The lag across the sensory signals consisted of relative phase shifts of the two square waves, while the raised cosine window remained constant (and synchronous across the senses). Audiovisual phase shift was manipulated according to the method of constant stimuli, and a full cycle was sampled in 40 steps. Each lag was presented 15 times to yield a total of 600 trials per participant. Besides the relative phase shifts across the multisensory signals, also the phase offset of the stimulus as a whole varied pseudorandomly across trials (spanning a full period sampled in 15 steps, one per repetition). Therefore, each time a given lag (phase shift) was tested; also the phase offset of the signals changed. Given that we expect the frequency doubling effect to be strong in size (i.e., synchrony at pi-phase shift should approach zero under the sustained hypothesis, and one under the transient hypothesis), and that we collected a large number of responses (n=600 per observer), a pool of five participants (age range 25–35 years, three females) was large enough to reliably assess its presence or lack thereof. Participants were paid 8 Euro/hour, and the experimental procedure was approved by the Ethics committee of the University of Bielefeld and was conducted in accordance with the declaration of Helsinki.

Results

Due to the periodic nature of the stimuli and hence of the experimental manipulation of lag, the resulting psychometric functions are also expected to be periodic, with an alternation of phase shifts yielding higher and lower reported synchrony. In this context, the evidence of frequency doubling can be measured from the number of oscillations in perceived simultaneity for a full cycle of phase shifts between the audiovisual stimuli.

Given the periodicity of the stimuli, it is natural to analyze the psychometric functions in the frequency domain. Therefore, to get a non-parametric and assumption-free estimate of the frequency of oscillations in the data, we ran a Fourier analysis on the empirical psychometric curves. If human responses are driven by transient input channels, we predict a peak at 2 cycles-per-period (cpp. i.e., frequency doubling), otherwise, there should be a peak at 1 cpp. The power spectrum of the psychometric functions shows a sharp peak at a frequency of 2 cpp in all observers, thereby indicating the presence of the hypothesized frequency doubling effect (**Figure 3B** and **Supplementary Figure S2**). To assess whether the amplitude at 1 and 2 cpp are statistically different at the individual observer level, we used a bootstrap procedure to estimate the confidence intervals of the response spectrum. For that, we used the binomial distribution and simulated 50000 psychometric functions, on which we performed a frequency analysis to obtain the 99% confidence intervals of the amplitudes. The results for both the aggregate observer and the individual data) show a clear separation between the confidence intervals of the amplitude at

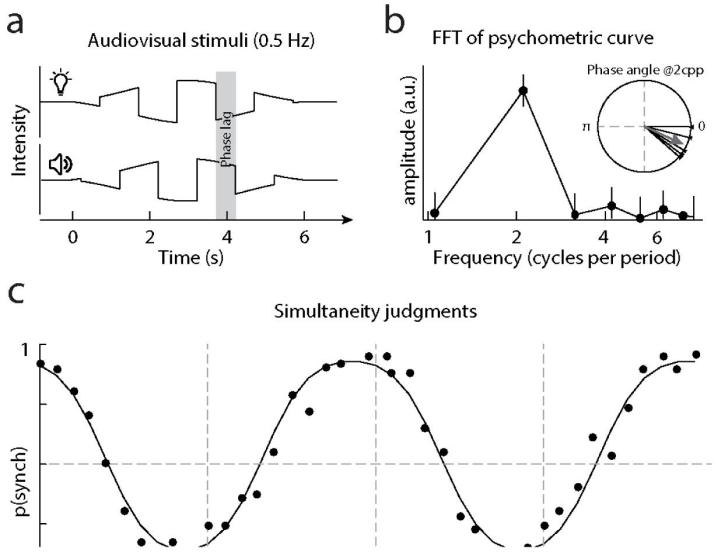


Figure 3.

Experiment 2, stimuli and results.

A. Schematic representation of the periodic stimuli. **B.** Frequency domain representation of the psychometric function: note the amplitude peak at 2 cycles per period. Errorbars represent the 99% confidence intervals. The inset represents the phase angle of the 2 cycles per period frequency component for each participant (thin lines), and the average phase (arrow). **C.** Results of Experiment 2 and psychometric fit, averaged across all participants. Each dot corresponds to 75 trials.

the frequencies of 1 and 2 cpp and demonstrate that the 2 cpp is indeed the dominant frequency, whereas the amplitude at 1 cpp is close to 0 and it is no different from the background noise (i.e., higher harmonics; see **Figure 3B**, and **Supplementary Figure S2**).

Given that the frequency analyses revealed a 2 cpp peak for all participants, we used this information to fit a sinusoidal psychometric function to the psychophysical data. Under the assumption of late Gaussian noise, the psychometric function can be written as:

$$p(synch|\phi) = normcdf(\alpha + \beta \cdot sin(f\phi + \theta)),$$

where the parameter α is the bias term, β the sensitivity, φ the phase-lag (range=[0, 2π]), f the frequency of oscillations, and θ is the phase offset (this shifts peak synchrony toward either positive or negative phases). Based on the Fourier analyses, we fixed the frequency (*f*) to 2 cpp, θ to -0.441 radians, and kept the linear coefficients–α and β–as free parameters, which were fitted to the psychophysical data using an adaptive Bayesian algorithm (L Acerbi 2017).

Overall, the fitted psychometric curves match the empirical data with a high goodness of fit (median r-squared=0.9321). Such an agreement between psychometric curves and empirical data achieved with the frequency parameter constrained by the Fourier analyses further indicates a reliable frequency doubling effect in all our participants.

Discussion

The results of Experiment 2 clearly demonstrate the existence of a frequency-doubling effect in the perceived simultaneity of periodic audiovisual stimuli. Given that the frequency doubling effect is only expected to occur if audiovisual synchrony is computed over unsigned transient input channels, the present results support a dominance of transient input channels in multisensory time perception. These results complement the conclusions of Experiment 1 and together demonstrate the key role of transient input channels in audiovisual integration.

Modelling

To account for multisensory integration, we have previously proposed a computational model, the Multisensory Correlation Detector (MCD, Parise and Ernst 2016), that exploits the temporal correlation between the senses to solve the correspondence problem, detect simultaneity and lag across the senses, and perform Bayesian-optimal multisensory integration. Based on the Hassenstein-Reichardt motion detector (Hassenstein and Reichardt 1956; Fujisaki and Nishida 2005, 2007), the core of the MCD is composed of two mirror-symmetric subunits, each multiplying visual and auditory input after applying a low-pass filter to each of them. As a consequence of this asymmetric filtering, each subunit is selectively tuned to different temporal order of the signals (that is, vision vs. audition lead). The outputs of the two subunits are then combined in different ways to detect the correlation and lag of multisensory signals, respectively. Specifically, correlation is calculated by multiplying the outputs of the subunits, hence producing an output (MCD_{Corr}) whose magnitude represents the correlation between the signals (**Figure 4C**). Temporal lag is instead detected by subtracting the outputs of the subunits, like in the classic Hassenstein–Reichardt detector (Hassenstein and Reichardt 1956). This yields an output (MCD_{Lag}) with a sign that represents the temporal order of the signals (**Fig. 4B**). While a single MCD unit can only perform temporal integration of multisensory input, a population of MCD units, each receiving input from spatially-tuned receptive fields (**Figure 4D**), followed by divisive normalization (**Figure 4E**, see Ohshiro, Angelaki, and DeAngelis 2011), can perform Bayesian-optimal spatial cue integration (e.g., see Alais and Burr 2004) for audiovisual source localization (**Figure 4F**, see Parise and Ernst 2016, for details)

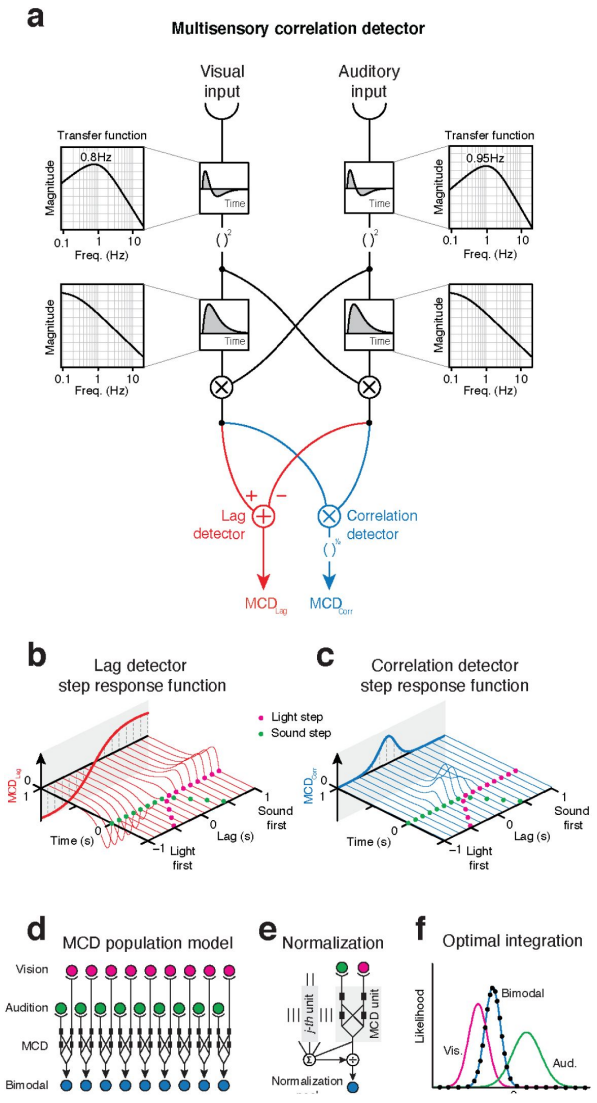


Figure 4.

MCD model.

A. Model schematics: the impulse-response functions of the channels are represented in the small boxes, and call-outs represent the transfer functions. **B.** Lag detector step responses as a function of the lag between visual and acoustic steps. **C.** Correlation detector responses as a function of the lag between visual and acoustic steps. **D.** Population of MCD units, each receiving input from spatiotopic receptive fields. **E.** Normalization, where the output of each unit is divided by the sum of the activity of all units. **F.** Optimal integration of audiovisual spatial cues, as achieved using a population of MCDs with divisive normalization. Lines represent the likelihood functions for the unimodal and bimodal stimuli; dots represent the response of the MCD model, which is indistinguishable from the bimodal likelihood function.

In its original form, the input to the MCD consisted of sustained unimodal channels, modelled as low-pass temporal filters (see also, (Burr et al. 2009; Yarrow et al. 2011). Although such a model could successfully account for the integration of trains of audiovisual impulses, the MCD cannot replicate the present results: for that, we first need to feed the model with unsigned transient unimodal input channels. Therefore, we replaced the front-end low-pass filters with band-pass temporal filters (to detect transients) followed by a squaring non-linearity (to get the unsigned transient (Stigliani, Jeska, and Grill-Spector 2017), and tested the model against the results of both Experiments 1 and 2, plus a variety of previously published psychophysical studies.

The equations of the revised MCD model are reported in the Supplementary Information. Just like the original version, the revised MCD model has three free parameters, representing the time constants of the two band-pass filters (one per modality) and that of the low-pass temporal filters of the subunits of the detector. Given that the tuning of the time constants of the model depends on the temporal profile of the input stimuli (see (Pesnot Lerousseau et al. 2022), here we used the data from Experiments 1 and 2 to constrain the temporal constants for the simulation of datasets consisting of either step stimuli or variations thereof (e.g., periodic stimuli), while for the simulation of experiments relying on stimuli with faster temporal rates, we constrained the temporal constants using data from Parise and Ernst (Parise and Ernst 2016). Details on parameter fitting are reported in the methods section. Given that in a previous study (Parise and Ernst 2016) we have already shown that alternative models for audiovisual integration are not flexible enough to reproduce the wide gamut of psychophysical data successfully accounted for by our model, here we only consider the MCD model (in its revised form). A Matlab script with the full implementation of the revised MCD model will be made publicly available before publication.

Simulation of Experiments 1 and 2

To test whether an MCD that receives input from unimodal transient channels could replicate the results of Experiments 1 and 2, we fed the stimuli to the detector and used the $\underline{MCD_{Corr}}$ (Equation 8) output to model simultaneity judgments, and the MCD_{Lag} (Equation 9) for the temporal order judgments (see Parise and Ernst 2016. Given that the psychophysical data consisted of response probabilities (i.e., probability of "synchronous" responses for the SJ and probability of "audio first" responses in the TOJ), whereas the outputs of the model are continuous variables expressed in arbitrary units, we used a GLM with a probit link function to transform the output of the MCD into probabilities (**Figure 5A**, see also (Parise and Ernst 2016), for a detailed description of the approach). The linear coefficients (i.e., slope and intercept) were fitted separately for each condition and task so that each psychometric curve had two free parameters for the GLM. The three parameters defining the temporal constants of the MCD model, instead, were fitted using all data from Experiments 1 and 2 combined using an adaptive Bayesian algorithm (L Acerbi 2017).

Overall, the model could tightly replicate the results of our experiments (**Figure 5B-D**): the Pearson correlation between the psychophysical data and model responses, computed across all conditions and participants was 0.99 for the SJ task of Experiment 1, 0.99 for the TOJ task of Experiment 1, and 0.97 for Experiment 2 (**Figure 5E**, see Supplementary Table 1). For comparison, the correlation between the data and the psychometric fits for Experiment 1 (i.e., cumulative Gaussians for the TOJs and difference of two cumulative Gaussians for the SJ) was 0.97, and 0.98 for Experiment 2; however, the psychometric fits required nearly twice as many free parameters compared to the model fits, and do not account for the generative process that give rise to the observed data. Importantly, just like human responses, the responses of the revised MCD model were nearly identical across all conditions in Experiment 1, whereas in Experiment 2 they displayed a clear frequency doubling effect. Furthermore, unlike the psychometric fits, which require to specify a priori the shape of the fitting function (e.g., sigmoid, bell, or periodic), the MCD model provides an output without specifying a priori any shape for the psychometric function. As shown in **Figure 5B-D** the model naturally captures the shape of the psychometric data. For instance, the same MCD_{Corr} output could generate bell-shaped response distributions in

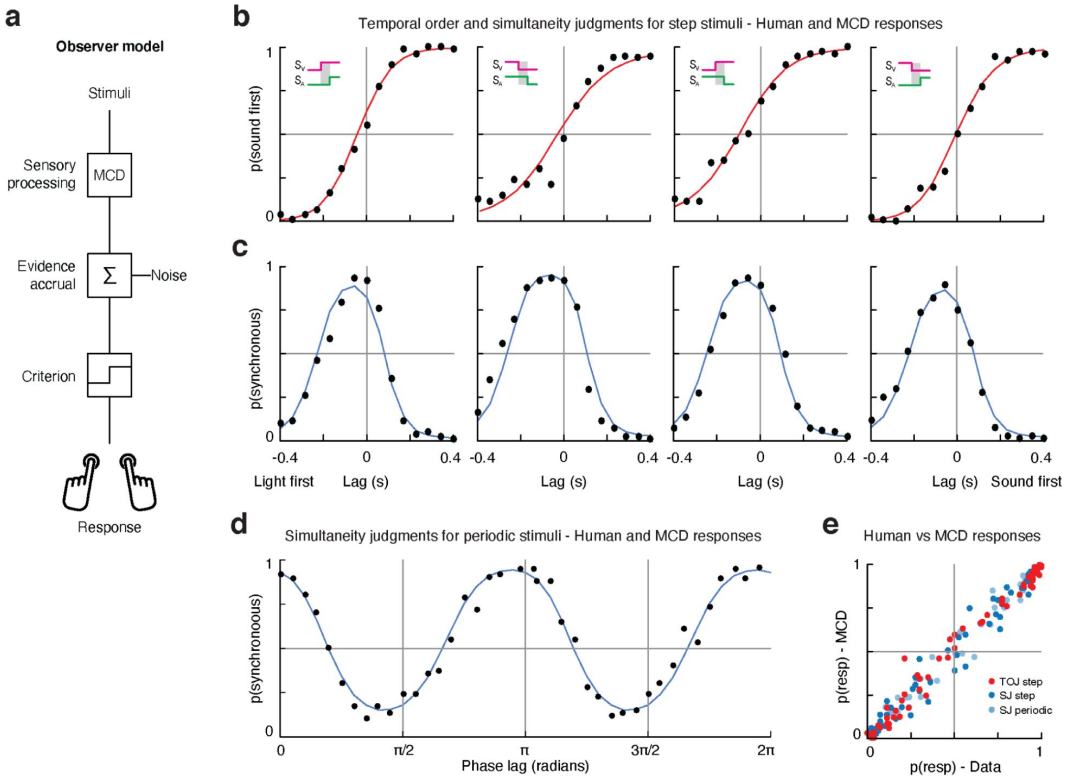


Figure 5.

MCD simulations of Experiments 1 and 2.

A. Schematics of the observer model, which receives input from one MCD unit to generate a behavioural response. The output of the MCD unit is integrated over a given temporal window (whose width depends on the duration of the stimuli), and corrupted by late additive noise before being compared to an internal criterion to generate a binary response. Such a perceptual decision-making process is modelled using a generalised linear model (GLM), depending on the task, the predictors of the GLM were either $MCD_{Corr.}$ (Equation 8) or MCD_{Lag} (Equation 9). **B.** Responses for the TOJ task of Experiment 1 (dots) and model responses (red curves). **C.** Responses for the SJ task of Experiment 1 (dots) and model responses (blue curves). **D.** Experiment 2 human (dots) and model responses (blue curve). **E.** Scatterplot of human vs model responses for both experiments.

Experiment 1 and sinusoidal responses in Experiment 2, purely based on the features of the input signals (such as its periodicity, or lack thereof). Therefore, taken together, the present simulations demonstrate that multisensory perception does indeed operate on correlated input from unimodal transient channels.

Validation of the MCD through simulation of published results

Although the simulations of Experiments 1 and 2 demonstrate that an MCD unit that receives input from transient unimodal channels is capable of reproducing the present psychophysical results, it is important to assess the generalizability of this approach. Therefore, we relied on the previous literature on multisensory perception of correlation, simultaneity and lag, as well as on the performance on crossmodal detection tasks, to validate our computational framework, and to assess its generalizability (by comparing the MCD responses against human performance on tasks not originally designed around our model). To this end, we selected a series of studies that employed parametric manipulations of the temporal structure of the signals, we simulated the stimuli, and we used the MCD model to predict human performance. If an MCD unit that receives input from transient unimodal channels is indeed the elementary computational unit for multisensory temporal processing, we should be able to reproduce all of the earlier findings on multisensory perception with our revised MCD model. A summary of our simulations, listing the sample size, the number of observers, and the Pearson correlation of MCD and human responses, is reported in Table S1.

Causality and temporal order judgments for random sequences of audiovisual impulses

When we first proposed the MCD model (Parise and Ernst 2016), we tested it against a psychophysical experiment in which participants were presented with a random sequence of 5 clicks and 5 flashes over an interval of 1 second (**Supplementary Figure S3A**), and had to report whether the signals appeared to share a common cause (causality judgments) and which modality came first (temporal order judgment). To test whether the revised MCD model could also account for these previous findings, we fed the same stimuli to the model, and used the MCD_{Corr} (**Eq. 8**) output to simulate causality judgments, and MCD_{Lag} (**Eq. 9**) to simulate temporal order judgments. Given that the stimuli in this experiment had a much higher temporal rate than the stimuli used in Experiments 1 and 2, the temporal constants of the MCD were set as free parameters that we fitted to the experimental data. Due to the stochastic nature of the stimuli, we analysed the data using reverse correlation (**Supplementary Figure S3B**). The temporal constants of the MCD were fitted to maximize the Pearson correlation between the empirical and simulated responses, for both causality and temporal order judgments (for details on modeling and reverse correlation analyses, see **Supplementary Figure S3B** and Parise and Ernst 2016).

Overall, the model faithfully reproduced the experimental data, and the Pearson correlation between empirical and simulated classification images was 0.99 (**Figure 6A-B**). This result closely replicates the original simulations (Parise and Ernst 2016) performed with a version of the MCD that instead received input from sustained unimodal temporal channels (not transient as the current model). Given the differences between transient and sustained temporal channels, it may seem surprising that the two models are in such close agreement in the current simulation. However, it is important to remember that the stimuli in these experiments consisted of impulses (clicks and flashes), and the impulse response function of the sustained and transient channels as modeled in this study are indeed very similar. Therefore, such an agreement between the responses of the original and the revised versions of the MCD are indeed expected. Beside proving

that the revised model can replicate previous results, this simulation allowed us to estimate the temporal constants of the MCD for trains of clicks and flashes, which was necessary for the next simulations.

Causality judgments for random sequences of audiovisual impulses with high temporal rates

To identify the temporal features of the stimuli that promote audiovisual integration, Locke and Landy ((Locke and Landy 2017), Experiment 2) ran a psychophysical experiment that was very similar to the causality judgment task described in the previous section. Although the stimuli in the two studies were nearly identical, the temporal sequences used by Locke and Landy (Locke and Landy 2017) had a considerably higher temporal rate (range 8-14 impulses/s), longer duration (2s) and a more controlled temporal structure (**Supplementary Figure S4**). Participants observed the stimuli and performed a causality judgment task, whose results demonstrated that the perception of a common cause depended both on the correlation in the temporal structure of auditory and visual sequences (**Figure 5C**, left) and the maximum lag between individual clicks and flashes (**Figure 5C**, right).

Given the similarity of this study with the causality judgment of Parise and Ernst (2016), we followed the same logic described above to perform reverse correlation analyses (without smoothing the cross-correlograms). Moreover, the MCD simulations were performed with the same temporal constants used for the simulations of Parise and Ernst (Parise and Ernst 2016), so that this experiment is simulated with a fully-constrained model, with zero free parameters. Unlike our previous experiment, however, the stimuli used by Locke and Landy (Locke and Landy 2017) varied in temporal rate, and hence also the number of clicks and flashes differed across trials. Given that the MCD is sensitive to the total stimulus energy, we normalized the model responses by dividing the MCD_{Corr} (Eq. 8) output by the rate of the stimuli.

Reverse correlation analyses were performed at the single subject level using both participants and model responses (see (Parise and Ernst 2016), for details on how the continuous model output of the MCD was discretized into a dichotomous variable), and the average data is shown in **Figure 6C** (left panel). Overall, the MCD model could near-perfectly predict the empirical classification image (Pearson correlation>0.99). Besides the reverse correlation analyses, Locke and Landy (2017) measured how the maximum lag between individual clicks and flashes affected the perceived common cause of audiovisual sequences. The results, shown in **Figure 6C** (right panel), demonstrate that perception of a common cause decreased with increasing maximum audiovisual lag. Once again, the MCD model accurately predicts this finding (Pearson correlation=0.98) without any free parameters. Taken together, the present simulations demonstrate that the MCD model can account for the perception of a common cause between stochastic audiovisual sequences, even when the stimuli have a high temporal rate, and highlight the importance of both similarity in the temporal structure and crossmodal lag in audiovisual integration.

Temporal order judgment for onset and offset stimuli

To investigate the perceived timing of auditory and visual on- and offsets, Wen and colleagues (Wen et al. 2020), presented continuous audiovisual noise stimuli with step on- and offsets (**Supplementary Figure S5**). The authors parametrically manipulated the audiovisual lag between either of the onset or the offset of the audiovisual stimulus and asked participants to report the perceived temporal order of the corresponding on- or offset, respectively. Despite large variability across participants, the point of subjective simultaneity systematically differed across conditions, with acoustic stimuli more likely appearing to change before the visual stimuli in the onset as compared to the offset condition.

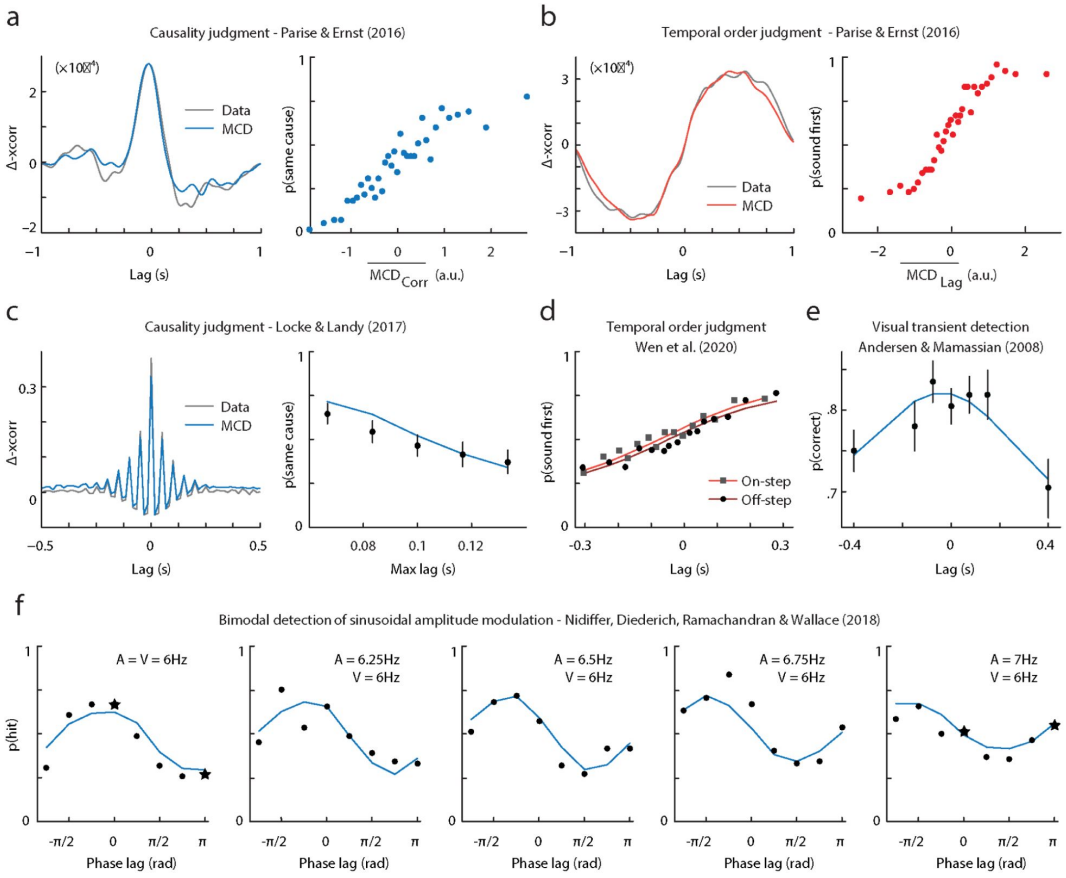


Figure 6.

MCD simulations of published results.

A. Results of the causality judgment task of Parise & Ernst (2016). The left panel represents the empirical classification image (grey) and the one obtained using the MCD model (blue). The right panel represents the output of the model plotted against human responses. Each dot corresponds to 315 responses. **B**. Results of the temporal order judgment task of Parise & Ernst (2016). The left panel represents the empirical classification image (grey) and the one obtained using the MCD model (red). The right panel represents the output of the model plotted against human responses. Each dot corresponds to 315 responses. **C**. Results of the causality judgment task of Locke & Landy (2017). The left panel represents the empirical classification image (grey) and the one obtained using the MCD model (blue). The right panel represents the effect of maximum audiovisual lag on perceived causality. Each dot represents on average 876 trials (range=[540, 1103]). **D**. Results of the temporal order judgment task of Wen et al. (2020). Squares represent the onset condition, whereas circles represent the offset condition. Each dot represents ≈745 trials. **E.** Results of the detection task of Andersen and Mamassian (2008), showing auditory facilitation of visual detection task. Each dot corresponds to 336 responses. **F.** Results of the audiovisual amplitude modulation detection task of Nidiffer and colleagues (2018), where the audiovisual correlation was manipulated by varying the frequency and phase of the modulation signals. Each dot represents ≈140 trials. The datapoint represented by a star correspond to the stimuli displayed in **Supplementary Figure S7**.

To test whether the model could replicate this finding, we generated audiovisual signals with the same temporal manipulations of audiovisual lag and fed them to the MCD. To minimize the number of free parameters, we ran the simulation using the temporal constants of the filters fitted using the data from Experiments 1 and 2 (see above). Moreover, to test whether the MCD alone could predict the PSS shift between the onset and offset condition, we combined the data from all participants, and used a single GLM to link MCD_{Lag} (**Eq. 9**) to the TOJ data, irrespective of condition. Therefore, this simulation consisted of just two free parameters (i.e., the slope and intercept of the GLM). Overall, the responses of the MCD were in excellent agreement with the empirical data (Pearson correlation=0.97) and successfully captured the difference across onset and offset condition (**Figure 6D**). For comparison, the Pearson correlation between the data and the psychometric functions (modeled as cumulative Gaussians, fitted independently for onset and offset conditions) was also 0.97, but required twice as many free parameters (two per condition), and it does not account for the underlying sensory information processing.

Acoustic facilitation of visual transients' detection

To study the temporal dynamic of audiovisual integration of unimodal transients, Andersen and Mamassian (Andersen and Mamassian 2008), Experiment 2) asked participants to detect a visual transient (i.e., a luminance increment) presented slightly before or after a task-irrelevant acoustic transient (i.e., a sound intensity increment, **Supplementary Figure S5**). The task consisted of a two-interval forced-choice, with the visual stimuli set at 75% detection threshold, and the asynchrony between visual and acoustic transients parametrically manipulated using the method of constant stimuli.

To simulate this experiment, we fed the stimuli to the MCD model, and used a GLM to link the MCD_{Corr} (**Eq. 8**) responses to the proportion of correct responses. The temporal constants of the model were constrained based on the results of Experiments 1 and 2, so that this simulation had two free parameters: the slope and the intercept of the GLM. Overall, the MCD could reproduce the data of Andersen and Mamassian (Andersen and Mamassian 2008), and the Pearson correlation between data and model responses was 0.91 (**Figure 6E**).

Detection of sinusoidal amplitude modulation

To test whether audiovisual amplitude modulation detection depends on the correlation between the senses, Nidiffer and colleagues ((Nidiffer et al. 2018), Experiment 2) asked participants to detect audiovisual amplitude modulation. Stimuli consisted of a pedestal intensity to which (in some trials) the authors added a near-threshold sinusoidal amplitude modulation (**Supplementary Figure S6A**). To manipulate audiovisual correlation, the authors varied the frequency and phase of the sinusoidal modulation signal. Specifically, the frequency of auditory amplitude modulation varied between 6Hz and 7Hz (5 steps), while the phase shift varied between 0 and 360 deg (8 steps). The frequency and phase shift of visual amplitude modulation were instead constant and set to 6Hz and 0 deg, respectively.

Both phase and frequency systematically affected participants' sensitivity; however, the results do not show any evidence of a frequency-doubling effect. This may appear as a surprising finding: given the analogy between this study and our Experiment 2, a frequency doubling effect should be intuitively expected (if, as we claim, correlation detection relied on transient input channels; though see **Supplementary Figure 7C**, for a visualization of the difference between these studies). In theory, when the amplitude modulations in the two modalities are 180 deg out of phase, the modulation signals are negatively correlated, and negatively correlated signals become positively correlated once fed to unsigned transient channels (**Supplementary Figure S6A**, bottom-left stimuli). When calculating the Pearson correlation of pairs of audiovisual signals, however, we should consider the whole stimuli: including the pedestal not just the amplitude modulations. Indeed, the stimuli used by Nidiffer and colleagues (2018) also consisted of linear ramps at onset and offset and a pedestal intensity level, compared with which the depth of the

amplitude modulation was barely noticeable (i.e., the modulation depth was about 6% of the pedestal level, see **Supplementary Figure 7C**). Therefore, once the audiovisual correlation is computed while also considering the ramps (and pedestals,see scatterplots in **Supplementary Figure S7A**), the lack of a frequency doubling effect becomes apparent: all stimuli used by Nidiffer and colleagues (2019) are strongly positively correlated, though such a correlation slightly varied across conditions (being 1 when amplitude modulation had the same frequency in both modalities and zero phase shift, and 0.8 in the least correlated condition, see **Supplementary Figure S6A-B**).

To simulate the experiment of Nidiffer and colleagues (Nidiffer et al. 2018), we fed the stimuli to the model and used the MCD_{Corr} (Eq. 8) output and a GLM to obtain the hit rate for the detection task. Given that neither the temporal constants of the MCD optimized for the Experiment 1 and 2, nor those optimized for Parise and Ernst (Parise and Ernst 2016), provided a good fit for Nidiffer's data, we set both the temporal constants of the MCD and the slope and intercept of the GLM as free parameters. All-in-all, the model could replicate the results of Nidiffer and colleagues (Nidiffer et al. 2018), and the Pearson correlation between the model and human responses was 0.89.

Conclusions

Taken together, the present results demonstrate the dominance of transient over sustained channels in audiovisual integration. Based on the assumption that perceived synchrony across the senses depends on the (Pearson) correlation of the unimodal signals (Wei 2006), we generated specific hypotheses for perceived audiovisual synchrony - depending on whether such a correlation is computed over transient or sustained inputs. Such hypotheses were then tested against the results of two novel psychophysical experiments, jointly showing that Pearson correlation between transient input signals systematically determines the perceived timing of audiovisual events. Based on that, we revised a general model for audiovisual integration, the Multisensory Correlation Detector (MCD), to selectively receive input from unimodal transient instead of sustained channels. Inspired by the motion detectors originally proposed for insect vision, such a biologically-plausible model integrates audiovisual signals through correlation detection, and could successfully account for the results of our psychophysical experiments along with a variety of recent findings in multisensory research. Specifically, once fed with transient input channels, the model could replicate human temporal order judgments, simultaneity judgments, causality judgments and crossmodal signal detection under a broad manipulation of input signals (i.e., step stimuli, trains of impulses, sinusoidal envelopes, etc.).

Previous research has already proposed a dominance of transient over sustained channels in audiovisual perception, with evidence coming from both psychophysical detection tasks and neuroimaging studies (Andersen and Mamassian 2008; Werner and Noppeney 2011). However, the role of transient and sustained channels on the perceived timing of audiovisual events (arguably the primary determinant of multisensory integration) has never been previously addressed, let alone computationally framed within a general model of multisensory integration. This study fills this obvious gap and explains all such previous findings in terms of the response dynamics of the MCD model, thereby comprehensively demonstrating that audiovisual integration relies on correlated input from unimodal transient channels. Recent research, however, criticized the MCD model for its alleged inability to either process audiovisual stimuli with a high temporal rate or detect signal correlation for short temporal intervals (Locke and Landy 2017; Colonius and Diederich 2020). By receiving inputs from transient unimodal channels, this revised version of the model fully addresses such criticisms. Indeed, the new MCD can near perfectly predict human performance in a causality judgment task with stimuli with a high temporal rate (up to 18Hz) using the same set of parameters optimized for stimuli with a much lower rate (5Hz), with a Pearson correlation coefficient of 0.99. Interestingly, recent results (Parise 2024) demonstrate

that a population of spatially tuned MCD units can also account for the integration of ecological audiovisual stimuli over time and space, thereby replicating phenomena such as the McGurk Illusion, the Ventriloquist Illusion and even attentional orienting.

While MCD simulations could replicate the shape of the empirical psychometric curves, standard psychometric functions can also fit the same datasets, sometimes with even higher goodness of fit. Hence one might wonder what the advantage of the current modelling approach is. However, when comparing MCD simulations with psychometric fits, it is important to focus on the differences between these two modelling approaches. Psychometric functions usually relate some physical parameter (the independent variable) to a measure of performance (the dependent variable)-based on statistical considerations regarding the stimuli and the underlying perceptual decision-making process. For example, based on assumptions on the nature of the bell-shaped distribution of audiovisual simultaneity judgments, Yarrow and colleagues proposed a model for simultaneity judgments that fits the SJ data of our Experiment 1 just as well as the MCD model (though with a larger number of free parameters). Psychometric curves for SJs, however, are not always bell-shaped: for example, in Experiment 2, they are sinusoidal. This finding is naturally captured by the MCD model but not by models that enforce some specific shape for the psychometric functions. The reason is that statistical approaches to psychometric curve fitting are usually agnostic as to how the raw input signals are transformed into evidence for perceptual decision-making. Indeed, while the inputs for psychometric fits are some parameters of the stimuli (e.g., the amount of audiovisual lag), the inputs for the MCD simulations are the actual stimuli themselves. That is the MCD extracts from the raw signals the evidence that is then fed into the perceptual decision-making process (i.e., the observer model, **Figure 5A**). By making explicit all the processing steps that link the input stimuli to a button press, the MCD model can tailor its predictions to any input stimuli, with the shape of the psychometric curves being unconstrained in principle, yet fully predictable based only on the input signals and the experimental task. This is why the same MCD model can predict bell-shaped psychometric functions in the SJ task of Experiment 1, sigmoidal functions in the TOJ task of Experiment 1, and sinusoidal functions in Experiment 2; whereas three different functions were necessary for the psychometric fits of the same experiments. Moreover, unlike psychometric fits, the MCD model is a biologically inspired neural model; as such it not only allows one to account for behavioural responses, but also for neurophysiological data, as recently shown through magnetoencephanlography (MEG, Pesnot Lerousseau et al. 2022)

Previous research attributed the multisensory benefits on perceptual decision-making tasks to “late”, post-sensory changes occurring at the level of the decision dynamics-with evidence stemming from EEG activity in brain regions commonly considered to be involved in “high-level”, decision-making (Franzen et al. 2020). While the present study cannot directly address such neurophysiological considerations, the computational framework proposed here, however, challenges any post-sensory interpretations. Indeed, the model proposed in this study can be divided into two separate components: the MCD, which handles the “early” sensory processing stage, and a “late” observer model, which receives input from the MCD and uses this information (along with the task demands) to generate a perceptual classification response (i.e., a button-press). In this context, it is important to note how multisensory benefits in detection tasks critically depend on the correlation and timing of audiovisual events: namely, the two factors that mostly affect the responses of the MCD model (e.g., see **Figure 5E-F**). Therefore, it is not surprising to see that the current framework can account for the effects of stimulus timing on performance purely based on the dynamics of the MCD responses, with no need for ad-hoc adjustments in the perceptual decision-making process. Further studies will hopefully reconcile such discrepant interpretations on the origins of multisensory benefits on perceptual decision-making tasks, possibly through a better understanding of the computations underlying the electrophysiological correlates recorded with modern imaging techniques.

Although the present study unequivocally supports a dominance of transient channels in multisensory integration, it is still necessary to consider which role, if any, sustained channels may play in crossmodal perception. Indeed, earlier studies have shown that sustained information, such as the intensity of visual and acoustic stimuli does indeed systematically affect performance in behavioural tasks (B. E. Stein et al. 1996; Odgaard, Arieh, and Marks 2004), and can even elicit phenomena known as crossmodal correspondences (C. Parise and Spence 2013). The mapping of intensity between vision and audition, however, is a somewhat peculiar one, as it is not obvious whether increasing intensity in one modality is mapped to higher or lower intensity in the other (that is, an obvious mapping between acoustic and lightness intensity, rooted in natural scene statistics, has not been reported, yet). A perhaps more profound mapping between sustained audiovisual information relates to redundant cues, that is, properties or a physical stimulus that can be jointly estimated via two or more senses; such as the size of an object, which can be simultaneously estimated through vision and touch (Ernst and Banks 2002, van Dam, Parise & Ernst, 2014). This is a particularly prominent aspect of multisensory perception, and the estimated size of an object is arguably sustained, rather than transient stimulus information. While the present study cannot directly address such a question, we propose that the correlation between visual and tactile transients, like those occurring when we reach and make contact with an object, is what the brain needs to solve the correspondence problem; and that hence let us infer that what we see and touch are indeed coming from the same distal stimulus. Then, once made sure that the spatial cues that we get from vision and touch are redundant, size information can be optimally integrated. Testing such a hypothesis is surely a fertile subject for future research.

Finally, it is important to consider what are the advantages of transient over sustained stimulus information for multisensory perception. An obvious one is parsimony, as dropping information that does not change over time entails lower transmission bandwidth by minimizing redundancy through efficient input coding (Barlow, 1961). Therefore, by only representing input variations, transient channels operate as event detectors, signalling the system of potentially relevant changes in the surrounding. Interestingly, a similar approach has grown in popularity in novel technological applications, such as neuromorphic circuits and event cameras. Indeed, while traditional sensors operate on a frame-based approach, whereby inputs are periodically sampled based on an internal clock, event cameras only respond to external changes in brightness as they occur, thereby reducing transmission bandwidth and maximizing dynamic range. Interestingly, recent work has even successfully exploited Hassenstein-Reichardt detectors as a biologically inspired solution for detecting motion with event cameras (D'Angelo et al. 2020). Hence, considering the mathematical equivalence of the MCD and the Hassenstein-Reichardt detectors, the present study suggests the intriguing possibility of using the revised MCD model as a biologically inspired solution for sensory fusion in future multimodal neuromorphic systems.

Supplementary information

The MCD model

The modified MCD model closely resembles the original model, but it takes input from unimodal transient (instead of sustained) input channels. That is, rather than being simply low-pass filtered, time-varying visual and auditory signals ($S_V(t)$, $S_A(t)$) are independently filtered by band-pass filters (f). Following Adelson and Bergen (1986), band-pass filters are modelled as biphasic impulse response functions defined as follows:

$$f_{mod}(t) = \frac{t}{\tau_{mod}} \cdot e^{\frac{-t}{\tau_{mod}}} \cdot \left(1 - \frac{t^2}{\tau_{mod}^2 \cdot n!}\right) \quad \text{(Equation 1).}$$

In this equation τ_{mod} is the modality-dependent temporal constant of the filter (mod=[a,v]). Based on the previous results of Experiments 1 and 2, we set these constants to be τ_V =0.070 s andτ_A =0.055 s for the visual and auditory filters, respectively. In line with Adelson and Bergen (1986), the parameter *n* (which controls the negative lobe of the impulse response) is set to 3.

As in the original implementation of the MCD model, the low-pass temporal filter of the correlation unit was:

$$f_{av}(t) = \frac{t}{\tau_{av}} \cdot e^{\frac{-t}{\tau_{av}}} \quad \text{(Equation 2)}$$

The temporal constant τ_{av} was set to 0.674 s. Although the bimodal temporal constant τ_{av} has a broad temporal tuning, the unimodal temporal constants τ_V and τ_A have systematic effects on the goodness of fit of the MCD model to our new psychophysical data. Therefore, in **Supplementary Figure S8** we show how the correlation between the MCD model and empirical data varies as a function of the temporal tuning of the visual and auditory transient detectors.

The filtered unisensory stimuli $Sf_{mod}(t)$ feeding into the correlation unit were obtained as follows:

$$Sf_{mod}(t) = [S_{mod}(t) * f_{mod}(t)]^2 \quad \text{(Equation 3)}$$

where (*) represents the convolution operator. Filtered signals are squared before being summed to render the responses insensitive to the polarity of changes in intensity (Stigliani, Jeska, and Grill-Spector 2017).

Like in the original MCD model, each sub-unit (u_1, u_2) of the detector independently combines filtered visual and auditory signals as follows:

$$u_1(t) = Sf_v(t) \cdot [Sf_a(t) * f_{av}(t)] \quad \text{(Equation 4),}$$

$$u_2(t) = Sf_a(t) \cdot [Sf_v(t) * f_{av}(t)] \quad \text{(Equation 5).}$$

To this end, the signals are convolved (*) with the low-pass temporal filters. The response of the sub-units is eventually multiplied or subtracted.

$$MCD_{corr}(t) = \sqrt{u_1(t) \cdot u_2(t)} \quad \text{(Equation 6),}$$

$$MCD_{corr}(t) = u_2(t) - u_1(t) \quad \text{(Equation 7).}$$

The resulting time-varying responses represent the local temporal correlation (MCD_{Corr}) and lag (MCD_{Lag}) across the signals. To reduce such time-varying responses into a single summary variable representing the total amount of evidence from each trial, we simply averaged the output of the detectors over a given temporal window of N samples:

$$\overline{MCD_{corr}} = \frac{1}{N}\sum_{t=1}^{N} MCD_{corr}(t) \quad \text{(Equation 8)},$$

$$\overline{MCD_{lag}} = \frac{1}{N}\sum_{t=1}^{N} MCD_{lag}(t) \quad \text{(Equation 9)}.$$

In the present simulations, the width of the temporal window varies across experiments, due to the variable duration of the audiovisual stimuli. The output of the MCD model is eventually transformed into probabilities using a general linear model with a probit link function (assuming additive Gaussian noise; see Parise & Ernst, 2016, for a similar approach).

Modelling sustained input channels

In line with previous work (Burr et al. 2009; Stigliani, Jeska, and Grill-Spector 2017; Parise and Ernst 2016), here we model sustained input channels as low-pass temporal filters with the following impulse response function:

$$f_{mod}(t) = \frac{t}{\tau_{mod}} \cdot e^{\frac{-t}{\tau_{mod}}} \quad \text{(Equation 10)}.$$

where τ_{mod} is the modality-dependent temporal constant of the filter (mod=[a,v]). Such a low-pass filter has the same shape as the bimodal filters of the MCD (Equation 2) and of the unimodal filter of the original MCD model (Parise and Ernst 2016).

Experiment 1: the relationship between the PSS and WSS measured using TOJs vs SJs

TOJs and SJs are the two main psychophysical tasks to measure sensitivity to lags across the senses. With both tasks it is possible to estimate the point and window of subjective simultaneity; however, when measured on the same subjects, the point and window of subjective simultaneity measured from the two tasks are often not correlated (García-Pérez and Alcalá-Quintana 2012; Linares and Holcombe 2014; Machulla, Di Luca, and Ernst 2016). This finding has been sometimes considered evidence for independent underlying neural mechanisms. Given that in Experiment 1, we estimated the PSS and WSS with both TOJs and SJs, we can repeat the same analyses on our dataset. Figure S1C shows the scatterplot of the PSS measured with the TOJ against the PSS measured with the SJ, and Figure S1D the scatterplot of the WSS measured with the TOJ against the SJ. Each point corresponds to one psychometric function in **Supplementary Figure S1A**. As in previous studies, the PSS and WSS measured with the two tasks are not significantly correlated (PSS: r=-0.16, p=0.38; WSS:r=0.16, p= 0.39).

While such a finding intuitively suggests the existence of independent mechanisms underlying the two tasks, our model clearly suggests otherwise. Indeed the MCD model provides two outputs: one representing the decision variable for the SJ (**Equation 8**), and the other for the TOJ (**Equation 9**). Therefore, we propose that TOJs and SJs share a common mechanism for sensory processing: the MCD model (**Equations 1**-**7**). However, the following decision-making processes (see **Figure 5A**) are independent across the two tasks, hence the lack of correlation between the PSS and WSS estimated using SJs vs TOJs.

Sample size

To properly substantiate our claims and quantitatively test our model, this study relies on a large collection of three novel psychophysical datasets, and six previously published ones. These consisted of behavioural responses from a variety of tasks such as temporal order judgments, simultaneity judgments, causality judgments and detection tasks, for a total of 68693 trials. Our Experiments 1 and 2 alone consisted of 12600 trials (4800 for the TOJ task in Experiment 1, 4800 for the SJ task in Experiment 1, and 3000 for Experiment 2). Given the nature of the present study, we were especially interested in determining the shapes of the psychometric functions, hence we prioritized collecting a large number of trials per observer (over a large pool of observers). Considering that we expected both large effect sizes (see **Figure 1B-C**) and low individual variability, the sample size of our new experiments is more than sufficient to draw reliable conclusions. Single subject analyses (**Supplementary Figures S1** and **S2**), showing consistent behaviour across participants support our original assumptions.

Nevertheless, it is important to stress that our psychophysical experiments only represent a small fraction of the overall dataset used in this study to assess and model the contribution of transient and sustained channels in multisensory integration. Indeed, when calculating the sample size of this study, we must also include all the previously published datasets that were re-analyzed and simulated with the MCD model. Hence, our conclusions are supported by a large-scale analysis and computational modelling of a vast set of behavioural data, consisting of 68693 trials from a sample of 110 observers, collectively providing strong converging evidence for the dominance of transient over sustained input channels in multisensory integration (See Table).

The data from Experiments 1 and 2 will be made publicly available online once the paper is accepted for publication.

The quadrature MCD model

The MCD units used so far receive input from unimodal transient channels modelled as a single biphasic temporal filter followed by squaring non linearities. In the original version of Adelson and Bergen (1986), however such a transient detection unit consisted of two biphasic temporal filters applied in parallel to the input, and the resulting signals are then squared and summed to each other. Although for the present simulations such a simplified version of the transient detector was sufficient, this may not be the case for different and more complex sets of stimuli. Hence, for completeness, here we also describe the full transient detector model.

Just like the simplified transient detector (**Equation 1** and **3**), also the full transient detector consists of biphasic temporal filters. However, instead of passing the signal through a single biphasic filter, the full transient detector consists of two biphasic temporal filters 90 degrees out of phase, applied in parallel to the incoming signals. Following Adelson and Bergen, these quadrature filters are modelled as follows:

$$f_n(t) = \left(\frac{t}{\tau_{bp}}\right)^n \bullet e^{-\frac{t}{\tau_{bp}}} \bullet \left[\frac{1}{n!} - \frac{1}{(n+2)!} \bullet \left(\frac{t}{\tau_{bp}}\right)^2\right] \qquad \text{Equation 11}$$

The phase of the filter is determined by n, which based on Emerson et al.[46] takes the values of 6 for the fast filter, and 9 for the slow one. The temporal constant of the filters is determined by the parameter τ_{bp}.

Fast and slow filters are applied to each unimodal input signal and the two resulting signals are squared and then summed. After that, a compressive non-linearity (square-root) is applied to the output, so as to constrain it within a reasonable range. Therefore, the output of each unimodal

unit feeding into the correlation detector takes the following form

$$Sf_{mod}(t) = \sqrt{\left[S_{mod}(t) * f_6(t)\right]^2 + \left[S_{mod}(t) * f_9(t)\right]^2} \quad \text{Equation 12}$$

where *mod* = *vid*, *aud* represents the sensory modality and * is the convolution operator.

These filtered signals are then multiplied in the two subunits of the MCD following the same logic as the simplified model (**Equations 4**-**9**).

Supplementary figure S8 displays how the temporal tuning of the unimodal temporal filter in the quadrature model affects the goodness of fit of the simulation of our Experiments 1 and 2. Clearly, the temporal tuning of the unimodal filters strongly affect the goodness of fit of the model: although overall a tuning below 0.06s is required for a good fit, the resulting landscape is highly irregular and overall favors slightly faster temporal constants for audition than vision. The parameter defining the low-pass temporal filter of the MCD subunits, instead, is less sensitive to the exact tuning. Hence for this figure its value is arbitrarily set to 0.674s. The psychometric functions generated by the full quadrature model are very similar to the ones generated by the simplified model (**Figure 5**).

Table S1.

Sample size of the datasets modelled and analysed in the present study.

Two of the datasets listed here (i.e., Experiment 1 and Parise et al., 2016) consisted of two tasks, each tested on the same pool of observers. The last row represents the total number of observers and trials, the average number of trials per participant, and the average correlation between MCD simulations and human data. Note how the revised MCD tightly replicated human responses in all of the datasets included in this study, despite major differences in stimuli, tasks, and sample sizes of the individual studies (see last column).

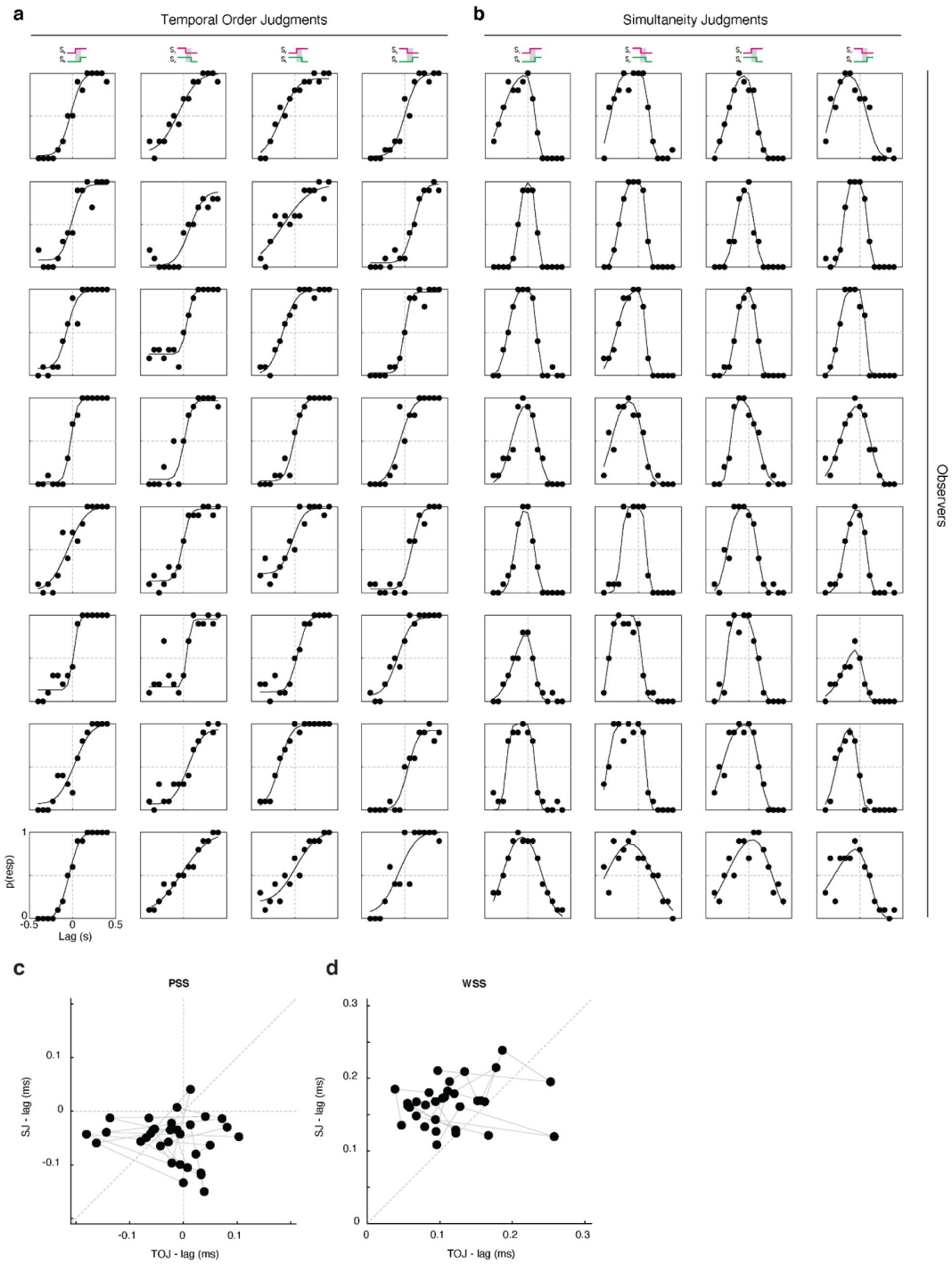
Dataset	Task	Number of observers	Number of trials	N. trials for observer	MCD-data Correlation
Experiment 1	Simultaneity judgment	8	4800	600	0.99
	Temporal order judgment		4800	600	0.99
Experiment 2	Simultaneity judgment	5	3000	600	0.98
Parise et al. (2016)	Causality judgment	5	9300	1860	0.98
	Temporal order judgment		9300	1860	0.99
Locke et al. (2017)	Causality judgment	10	7200	720	0.99
Wen et al. (2020)	Temporal order judgment	57	22337	≈392	0.97
Andersen et al (2008)	Transient detection	13	2352	≈181	0.91
Nidiffer et al. (2018)	Modulation detection	12	5604	467	0.89
	Overall	**110**	**68693**	**≈624**	**0.97**

Table S2.

Results of Friedman test and Bayesian Repeated Measures ANOVA for Experiment 1.

Four separate Friedman tests were used to assess whether the four experimental conditions differed in terms of PSS and WSS in the TOJ and SJ tasks. The first column represents the variables, the second the χ^2 value and the degrees of freedom (in brackets), the third the p-value. Given that we ran four tests on the same dataset, statistical significance should be computed by comparing the p-value against a Bonferroni-adjusted alpha level of 0.0125 (i.e., 0.05/4). The last column represents the Bayes factor BF_{01} in favour of the null hypothesis as calculated using Bayesian Repeated Measures ANOVA using the statistical software JASP (JASP Team 2024; Version 0.18.3) with the default settings. Together with the Friedman test, the present analyses provide further converging evidence for the lack of meaningful differences across the two tasks and four conditions of Experiment 1.

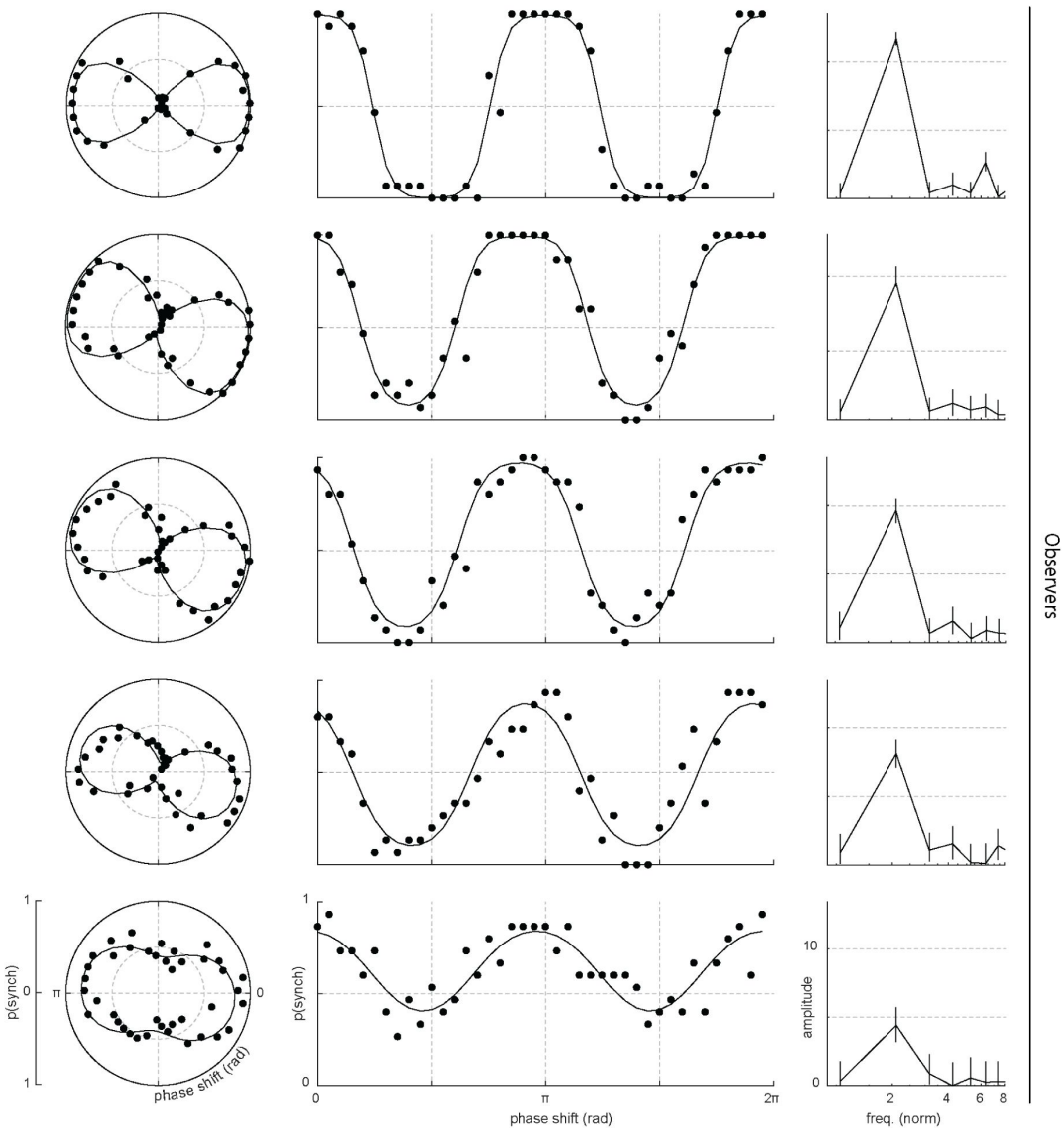
	Friedman test		Bayesian ANOVA
	χ^2(df)	p-value	BF_{01}
PSS (TOJ)	6.15 (3)	0.1045	2.833
PSS (SJ)	2.85 (3)	0.4153	2.237
WSS (TOJ)	2.55 (3)	0.4663	2.570
WSS (SJ)	5.55 (3)	0.1357	1.452



Supplementary Figure S1.

Results and psychometric fits of Experiment 1.

Data from different observers are represented in different rows. **A.** represent the data of the TOJ task, **B.** represent the data of the SJ task. The icons on top represent the different conditions. Each dot corresponds to 10 trials. **C.** represents the scatterplot of the PSS measured with the TOJ plotted against the PSS measured with the SJ. **D.** represents the scatterplot of the WSS measured with the TOJ plotted against the PSS measured with the SJ. Each dot in panels C and D represents the PSS or WSS from one condition and observer (that is, there are 4 dots for each observer, one per condition); datapoints from the same participant are linked by a grey line.



Supplementary Figure S2.

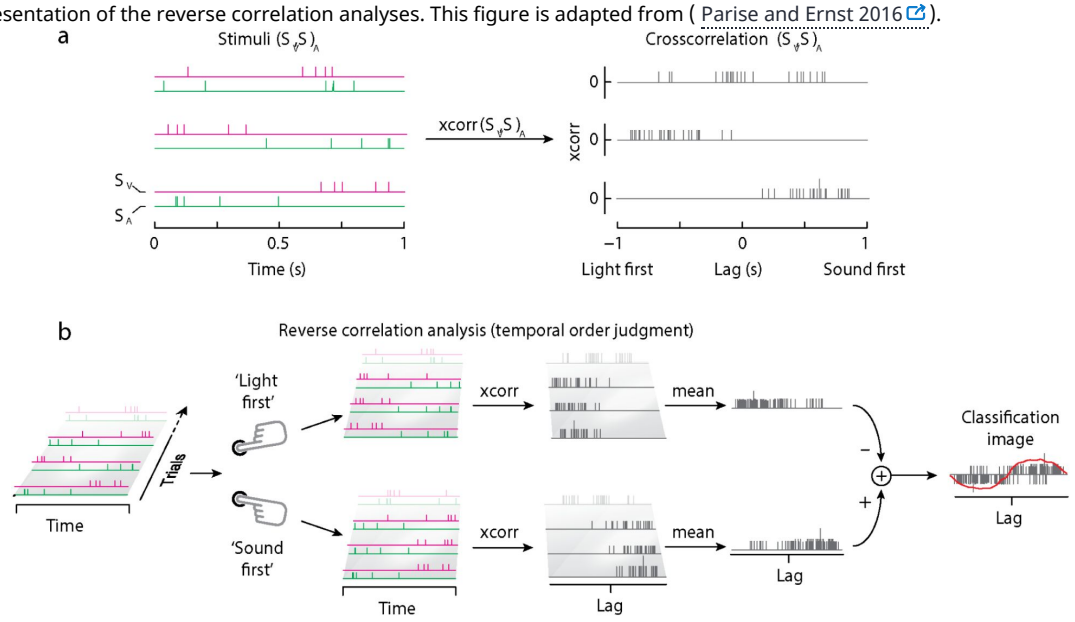
Results and psychometric fits of Experiment 2.

Data from different observers are represented in different rows. The first four column represent the data in polar coordinates, whereas the second column represent the same dataset in Cartesian coordinates. The counter-clockwise rotation of the polar psychometric functions indicate that maximum perceived synchrony across the senses occur when vision changes slightly before audition. Each dot corresponds to 15 trials. The last column represents the psychometric curve in the frequency domain: note the peak at 2cpp in every participant, indicating the frequency doubling effect. The errorbars in the last column represent the 99% confidence intervals.

Supplementary Figure S3.

Stimuli and reverse correlation analyses of Parise & Ernst (2016).

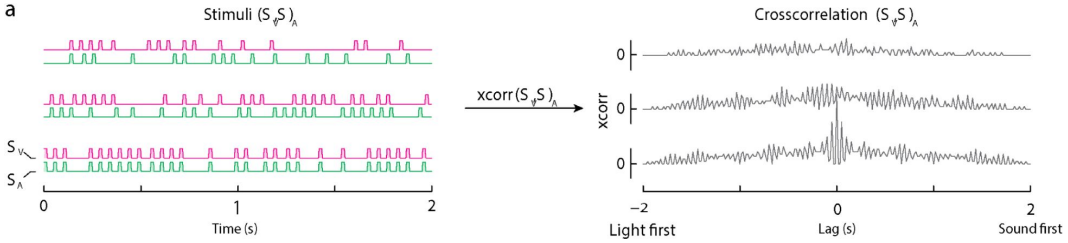
Panel **A** represents three pairs of audiovisual stimuli (left), and their cross-correlation (right). Panel **B** shows a schematic representation of the reverse correlation analyses. This figure is adapted from (Parise and Ernst 2016).



Supplementary Figure S4.

Stimuli of Locke & Landy (2017).

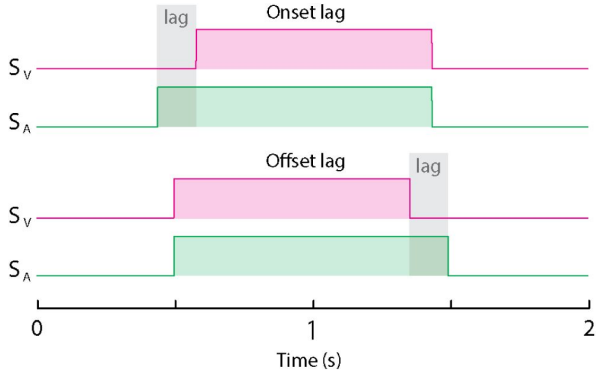
Three example pairs of audiovisual stimuli, and their cross-correlogram. Note how the stimuli vary in both terms of temporal rate (i.e., total number of clicks and flashes) and audiovisual correlation (with the bottom one being fully correlated).



Supplementary Figure S5.

Stimuli of Wen et al. (2020).

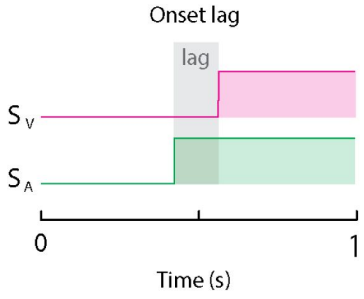
Stimuli consisted of rectangular temporal envelopes, with a variable audiovisual lag either at the onset (top) or offset (bottom). The amount of audiovisual lag varied across trials, following a staircase procedure.

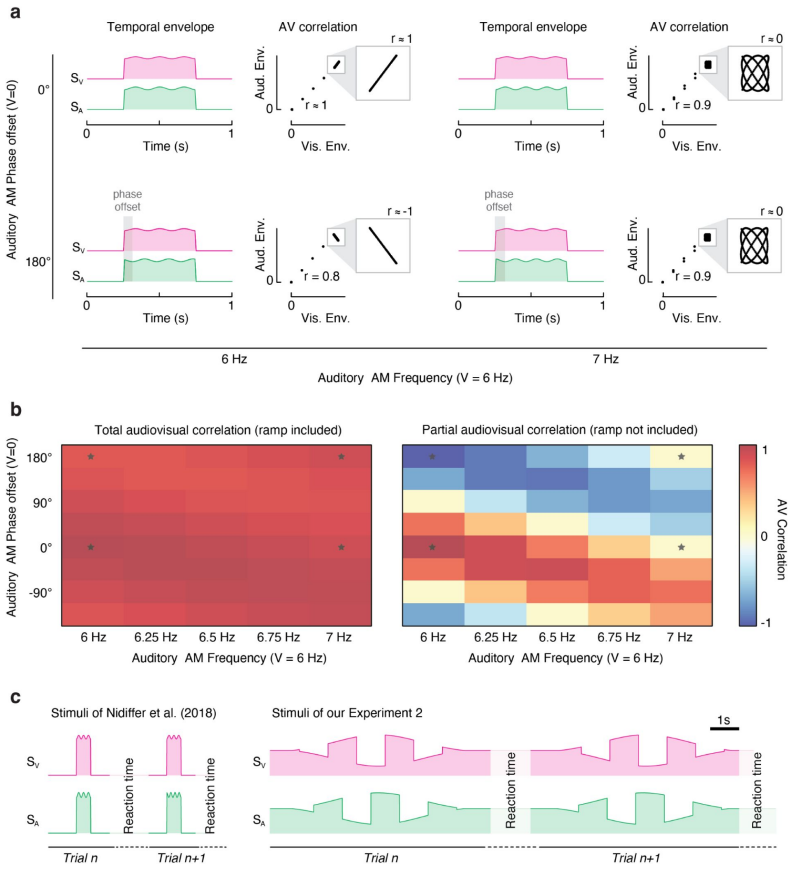


Supplementary Figure S6.

Stimuli of Andersen & Mamassian (2008).

Stimuli consisted of on-steps, with a parametrical manipulation of the lag between vision and audition, determined using the method of constant stimuli.

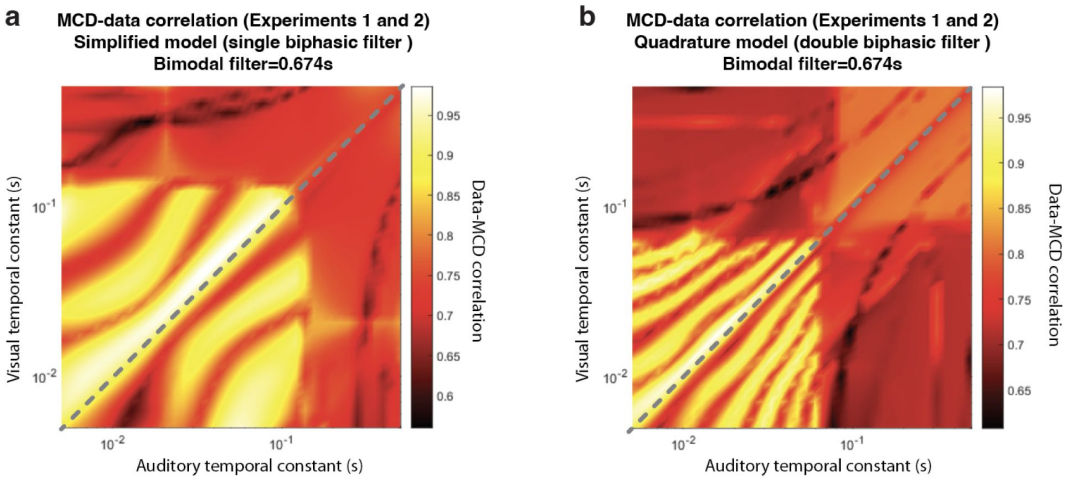




Supplementary Figure S7.

Stimuli of Nidiffer et al. (2018).

Stimuli consisted of a sinusoidal amplitude modulation over a pedestal intensity (**Panel A**). The pedestal had a 10ms linear ramp at onset and offset. While the visual sinusoidal amplitude modulation was constant throughout the experiment (frequency = 6Hz, zero phase offset), the auditory amplitude modulation varied in both frequency (from 6Hz to 7Hz, in 5 steps) and phase offset (from 0 to 360 deg, in 8 steps). Next to each pair of stimuli, we represent the scatterplot of the visual and auditory envelopes, which highlight the audiovisual correlation. The call-out boxes zoom in on the correlation between the audiovisual sinusoidal amplitude modulation (excluding the linear onset and offset ramps), whereas the main scatterplots also display the ramps. **Panel B** represents the audiovisual Pearson correlation for all the stimuli used by Nidiffer et al. The left panel shows the total audiovisual correlation, which was calculated while also including the onset and offset ramps; the right panel represents the partial correlation, calculated by only considering the sinusoidal amplitude modulation (i.e., the area inside the call-outs in Panel A) without considering the onset and offset linear ramps (as done in Nidiffer et al., 2018). Note that once the ramps are included in the analyses, all audiovisual stimuli are strongly correlated (with only minor differences across conditions). The stars in the correlation matrices mark the cells corresponding to the four stimuli represented in Panel A, and the datapoints represented by a star in **Figure 5F**. **Panel C** displays a comparison between the stimuli used by Nidiffer et al. (2018) and our Experiment 2. Both the time axis (abscissa) and the intensity envelope (ordinate) are here drawn to scale. Although both experiments consist of stimuli with periodic amplitude modulations, there are key important differences. First off, while between two consecutive trials the visual and auditory stimuli were completely off in Nidiffer, in our study the pedestal was always present (without any interruptions across consecutive trials). That introduces transients in Nidiffer's study both at the beginning and the end of each stimulus, which are absent in our stimuli. The relative magnitude of such transients to the comparatively low depth of amplitude modulation (of the sinusoidal component) is the ultimate reason for the absence of frequency doubling in Nidiffer et al. (2018). In our study, transients at the beginning and end of each trial are prevented by playing a constant pedestal stimulus level across trials, and by applying a Gaussian envelope to the depth of our square-wave modulation. Additionally, it is important to stress the obvious difference in the duration and frequency of the stimuli used in the two studies. Specifically, our stimuli were 12 times longer than the stimuli of Nidiffer, but the frequency of amplitude modulation of our study was about 12 times lower (varying depending on the various conditions of Nidiffer et al.). Finally, note the difference in the depth of amplitude modulation across the two experiments.



Supplementary Figure S8.

Effect of unimodal temporal constants on the goodness of fit (Pearson correlation) between the MCD model and the data from our Experiments 1 and 2. Panel A represents the simplified model, with a single biphasic temporal filter. Panel B represents the full model, with unimodal temporal filters in quadrature pairs.

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Editors

Reviewing Editor
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Senior Editor
Andrew King
University of Oxford, Oxford, United Kingdom

Reviewer #1 (Public Review):

The authors present a model for multisensory correlation detection that is based on the neurobiologically plausible Hassenstein Reichardt detector (Parise & Ernst, 2016). They demonstrate that this model can account for human behaviour in synchrony or temporal order judgements and related temporal tasks in two new data sets (acquired in this study) and a range of previous data sets. While the current study is limited to the model assessment for relatively simple audiovisual signals, in future communications, the authors demonstrate that the model can also account for audiovisual integration of complex naturalistic signals such as speech and music.

The significance of this work lies in its ability to explain multisensory perception using fundamental neural mechanisms previously identified in insect motion processing.

Strengths:

(1) The model goes beyond descriptive models such as cumulative Gaussians for TOJ and differences in cumulative Gaussians for SJ tasks by providing a mechanism that builds on the neurobiologically plausible Hassenstein-Reichardt detector.
(2) This model can account for results from two new experiments that focus on the detection of correlated transients and frequency doubling. The model also accounts for several behavioural results from experiments including stochastic sequences of A/V events and sine wave modulations (and naturalistic Av signals such as speech and music as shown in future communications).

https://doi.org/10.7554/eLife.90841.2.sa2

Reviewer #2 (Public Review):

Summary:

This is an interesting and well-written manuscript that seeks to detail performance on two human psychophysical experiments designed to look at the relative contributions of transient and sustained components of a multisensory (i.e., audiovisual) stimulus to their integration. The work is framed within the context of a model previously developed by the authors and now somewhat revised to better incorporate the experimental findings. The major takeaway from the paper is that transient signals carry the vast majority of the information related to the integration of auditory and visual cues, and that the Multisensory Correlation Detector (MCD) model not only captures the results of the current study, but is also highly effective in capturing the results of prior studies focused on temporal and causal judgments.

Strengths:

Overall the experimental design is sound and the analyses well performed. The extension of the MCD model to better capture transients make a great deal of sense in the current context, and it is very nice to see the model applied to a variety of previous studies.

Comments on the revised version:

In the revised manuscript, the authors have done an excellent job of responding to the prior critiques. I have no additional concerns or comments.

https://doi.org/10.7554/eLife.90841.2.sa1

Author response:

The following is the authors' response to the original reviews.

Public Reviews:

> ***Reviewer #1 (Public Review):***
>
> *The authors present a model for multisensory correlation detection that is based on the neurobiologically plausible Hassenstein Reichardt detector. It modifies their previously reported model (Parise & Ernst, 2016) in two ways: a bandpass (rather than lowpass) filter is initially applied and the filtered signals are then squared. The study shows that this model can account for synchrony judgement, temporal order judgement, etc in two new data sets (acquired in this study) and a range of previous data sets.*
>
> *Strengths:*
>
> *(1) The model goes beyond descriptive models such as cumulative Gaussians for TOJ and differences in cumulative Gaussians for SJ tasks by providing a mechanism that builds on the neurobiologically plausible Hassenstein-Reichardt detector.*
>
> *(2) This modified model can account for results from two new experiments that focus on the detection of correlated transients and frequency doubling. The model also accounts for several behavioural results from experiments including stochastic sequences of A/V events and sine wave modulations.*
>
> *Additional thoughts:*
>
> *(1) The model introduces two changes: bandpass filtering and squaring of the inputs. The authors emphasize that these changes allow the model to focus selectively on transient rather than sustained channels. But shouldn't the two changes be introduced separately? Transients may also be detected for signed signals.*

We updated the original model because our new psychophysical evidence demonstrates the fundamental role of unsigned transient for multisensory perception. While the original model received input from sustained unimodal channels (low-pass filters), the new version receives input from unsigned unimodal transient channels. Transient channels are normally modelled through bandpass filters (to remove the DC and high-frequency signal components) and squaring (to remove the sign). While these may appear as two separate changes in the model, they are, in fact, a single one: the substitution of sustained with unsigned transient channels (for a similar approach, see Stigliani et al. 2017, PNAS). Either change alone would not be sufficient to implement a transient channel that accounts for the present results.

That said, we were also concerned with introducing too many changes in the model at once. Indeed, we simply modelled the unimodal transient channels as a single band-pass filter

followed by squaring. This is already a stripped-down version of the unsigned transient detectors proposed by Adelson and Bergen in their classic Motion Energy model. The original model consisted of two biphasic temporal filters 90 degrees out of phase (i.e., quadrature filters), whose output is later combined. While a simpler implementation of the transient channels was sufficient in the present study, the full model may be necessary for other classes of stimuli (including speech, Parise, 2024, BiorXiv). Therefore, for completeness, we now include in the Supplementary Information a formal description of the full model, and validate it by simulating our two novel psychophysical studies. See Supplementary Information "The quadrature MCD model" section and Supplementary Figure S8.

> *(2) Because the model is applied only to rather simple artificial signals, it remains unclear to what extent it can account for AV correlation detection for naturalistic signals. In particular, speech appears to rely on correlation detection of signed signals. Can this modified model account for SJ or TOJ judgments for naturalistic signals?*

It can. In a recent series of studies we have demonstrated that a population of spatially-tuned MCD units can account for audiovisual correlation detection for naturalistic stimuli, including speech (e.g. the McGurk Illusion). Once again, unsigned transients were sufficient to replicate a variety of previous findings. We have now extended the discussion to cover this recent research: Parise, C. V. (2024). Spatiotemporal models for multisensory integration. bioRxiv, 2023-12.

> *Even Nidiffer et al. (2018) which is explicitly modelled by the authors report a significant difference in performance for correlated and anti-correlated signals. This seems to disagree with the results of study 1 reported in the current paper and the model's predictions. How can these contradicting results be explained? If the brain detects correlation on signed and unsigned signals, is a more complex mechanism needed to arbitrate between those two?*

We believe the reviewer here refers to our Experiment 2 (where, like Nidiffer at al. (2018) we used periodic stimuli, not Experiment 1, which consists of step stimuli). We were also puzzled by the difference between our Experiment 2 and Nidiffer et al. (2018): we induced frequency doubling, Nidiffer did not. Based on quantitative simulations, we concluded that this difference could be attributed to the fact that while Nidiffer included on each trial an intensity ramp in their periodic audiovisual stimuli, we did not. As a result, when considering the ramp (unlike in Nidiffer's analyses), all audiovisual signals used by Nidiffer were positively correlated (irrespective of frequency and phase offset), while our signals in Experiment 2 were sometimes correlated and other times not (depending on the phase offset). This important simulation is included in Supplementary Figure S7; we also have now updated the text to better highlight the role of the pedestal in determining the direction of the correlation.

> *(3) The number of parameters seems quite comparable for the authors' model and descriptive models (e.g. PSF models). This is because time constants require refitting (at least for some experimental data sets) and the correlation values need to be passed through a response mode (i.e. probit function) to account for behavioural data. It remains unclear how the brain adjusts the time constants to different sensory signals.*

This is a deep question. For simplicity, here the temporal constants were fitted to the empirical psychometric functions. To avoid overfitting, whenever possible we fitted such parameters over some training datasets, while trying to predict others. However, in some cases, it was necessary to fit the temporal constants to specific datasets. This may suggest that the temporal tuning of those units is not crystalised to some pre-defined values, but is

adjusted based on recent perceptual history (e.g., the sequence of trials and stimuli participants are exposed to during the various experiments).

For transparency, here we show how varying the tuning of the temporal constants of the filters affects the goodness of fit of our new psychophysical experiments (Supplementary Figure S8). As it can be readily appreciated, the relative temporal tuning of the unimodal transient detector was critical, though their absolute values could vary over a range of about 15 to over 100ms. The tuning of the low-pass filters of the correlation detector (not shown here) displayed much lower temporal sensitivity over a range between 0.1s to over 1s.

This simulation shows the impact of temporal tuning in our simulations, however, the question remains as to how such a tuning gets selected in the first place. An appealing explanation relies on natural scene statistics: units are temporally tuned to the most common audiovisual stimuli. Although our current empirical evidence does not allow us to quantitatively address this question, in previous simulations (see Parise & Ernst, 2016, Supplementary Figure 8), by analogy with visual motion adaptation, we show how the temporal constants of our model can dynamically adjust and adapt to recent perceptual history. We hope these new and previous simulations address the question about the nature of the temporal tuning of the MCD units.

> *(4) Fujisaki and Nishida (2005, 2006) proposed mechanisms for AV correlation detection based on the Hassenstein-Reichardt motion detector (though not formalized as a computational model).*

This is correct, Fujisaki and Nishida (2005, 2007) also hypothesized that AV synchrony could be detected using a mechanism analogous to motion detection. Interestingly, however, they ruled out such a hypothesis, as their "data do not support the existence of specialized low-level audio-visual synchrony detectors". Yet, along with our previous work (Parise & Ernst, 2016, where we explicitly modelled the experiments of Fujisaki and Nishida), the present simulations quantitatively demonstrate that a low-level AV synchrony detector is instead sufficient to account for audiovisual synchrony perception and correlation detection. We now credit Fujusaki and Nishida in the modelling section for proposing that AV synchrony can be detected by a cross-correlator.

Finally, we believe the reviewer is referring to the 2005 and 2007 studies of Fujisaki and Nishida (not 2006); here are the full references of the two articles we are referring to:

Fujisaki, W., & Nishida, S. Y. (2005). Temporal frequency characteristics of synchrony–asynchrony discrimination of audio-visual signals. Experimental Brain Research, 166, 455-464.

Fujisaki, W., & Nishida, S. Y. (2007). Feature-based processing of audio-visual synchrony perception revealed by random pulse trains. Vision Research, 47(8), 1075-1093.

> ***Reviewer #2 (Public Review):***
>
> *Summary:*
>
> *This is an interesting and well-written manuscript that seeks to detail the performance of two human psychophysical experiments designed to look at the relative contributions of transient and sustained components of a multisensory (i.e., audiovisual) stimulus to their integration. The work is framed within the context of a model previously developed by the authors and is now somewhat revised to better incorporate the experimental findings. The major takeaway from the paper is that transient signals carry the vast majority of the information related to the integration of auditory and visual cues, and that the Multisensory Correlation Detector (MCD) model not only captures the results of*

> *the current study but is also highly effective in capturing the results of prior studies focused on temporal and causal judgments.*
>
> *Strengths:*
>
> *Overall the experimental design is sound and the analyses are well performed. The extension of the MCD model to better capture transients makes a great deal of sense in the current context, and it is very nice to see the model applied to a variety of previous studies.*
>
> *Weaknesses:*
>
> *My one major issue with the paper revolves around its significance. In the context of a temporal task(s), is it in any way surprising that the important information is carried by stimulus transients? Stated a bit differently, isn't all of the important information needed to solve the task embedded in the temporal dimension? I think the authors need to better address this issue to punch up the significance of their work.*

In hindsight, it may appear unsurprising that transient signals carry most information for audiovisual integration. Yet, so somewhat unexpectedly, this has never been investigated using perhaps the most diagnostic psychophysical tools for perceived crossmodal timing; namely temporal order and simultaneity judgments–along with carefully designed experiments with quantitative predictions for the effect of either channel. The fact that the results conform to intuitive expectations further supports the value of the present work: grounding empirically with what is intuitively expected. This offers solid psychophysical evidence that one can build on for future advancements. Importantly, developing a model that builds on our new results and uses the same parameters to predict a variety of classic experiments in the field, further supports the current approach.

If “significance” is intended as shaking previous intuitions or theories, then no: this is not a significant contribution. If instead, by significance we intend to build a solid empirical and theoretical ground for future work, then we believe this study is not significant, it is foundational. We hope that this work's significance is better captured in our discussion.

On a side note, there is an intriguing factor around transient vs. sustained channels: what matters is the amount of change, not the absolute stimulus intensity. Previous studies, for example, have suggested a positive cross modal mapping between auditory loudness and visual lightness or brightness [Odegaard et al., 2004]. This study, conversely, challenges this view and demonstrates that what matters for multisensory integration in time is not the intensity of a stimulus, but changes thereof.

> *In a more minor comment, I think there also needs to be a bit more effort into articulating the biological plausibility/potential instantiations of this sustained versus transient dichotomy. As written, the paper suggests that these are different "channels" in sensory systems, when in reality many neurons (and neural circuits) carry both on the same lines.*

The reviewer is right, in our original manuscript we glossed over this aspect. We have now expanded the introduction to discuss their anatomical basis. However, we are not assuming any strict dichotomy between transient and sustained channels; rather, our results and simulations demonstrate that transient information is sufficient to account for audiovisual temporal integration.

> ***Recommendations for the authors:***

> ***Reviewer #1 (Recommendations For The Authors):***
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> *(1) Related to point 2 of the public review, can the authors provide additional results showing that the model can also account for naturalistic signals and more complex stochastic signals?*

While working on this manuscript, we were also working in parallel on a project related to audiovisual integration of naturalistic signals. A pre-print is available online [Parise, 2024, BiorXiv], and the related study is now discussed in the conclusions.

> *(2) As noted in the public review, Fujisaki and Nishida (2005, 2006) already proposed mechanisms for AV correlation detection based on the Hassenstein-Reichardt motion detector. Their work should be referenced and discussed.*

We have now acknowledged the contribution of Fujisaki and Nishida in the modelling section, when we first introduce the link between our model and the Hassenstein-Reichardt detectors.

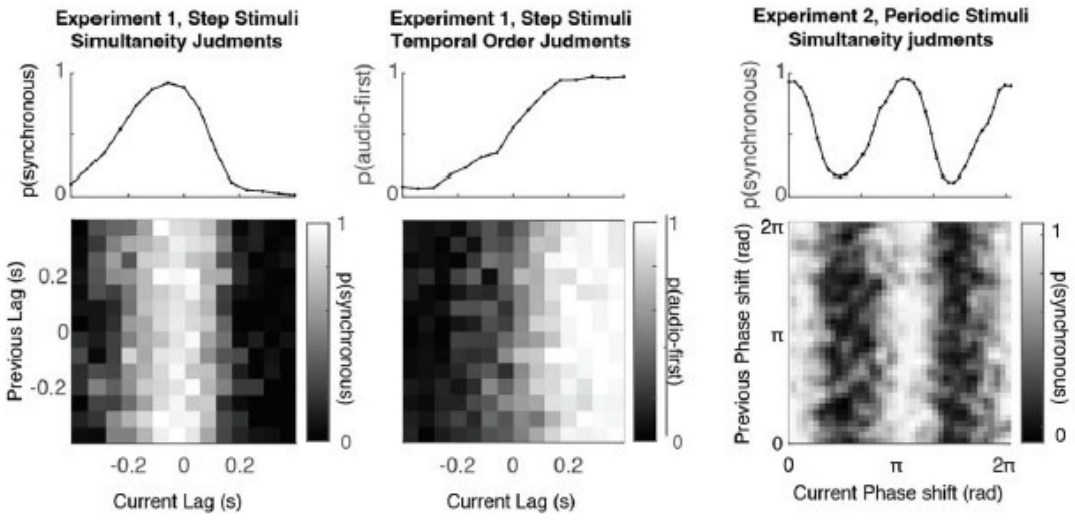
> *(3) Experimental parameters: Was the phase shift manipulated in blocks? If yes, what about temporal recalibration?*

To minimise the effect of temporal recalibration, the order of trials in our experiments was randomised. Nonetheless, we can directly assess potential short-term recalibration effects by plotting our psychophysical responses against both the current SOA, and that of the previous trials. The resulting (raw) psychometric surfaces below are averaged across observers (and conditions for Experiment 1). In all our experiments, responses are obviously dependent on the current SOA (x-axis). However, the SOA of the previous trials (y-axis) does not seem to meaningfully affect simultaneity and temporal order judgments. The psychometric curves above the heatmaps represent the average psychometric functions (marginalized over the SOA of the previous trial).

All in all, the present analyses demonstrate negligible temporal recalibration across trials, likely induced by a random sequence of lags or phase shifts. Therefore, when estimating the temporal constants of the model, it seems reasonable to ignore the potential effects of temporal recalibration. To avoid increasing the complexity of the present manuscript, we would prefer not to include the present analyses in the revised version.

Author response image 1.

Effect of previous trial. Psychometric surfaces for Experiments 1 and 2 plotted against the lag in the current vs. the previous trial. While psychophysical responses are strongly modulated by the lag in the last trial (horizontal axis), they are relatively unaffected by the lag in the previous trial (vertical axis).



> *(4) The model predicts no differences for experiment 1 and this is what is empirically observed. Can the authors support these null results with Bayes factors?*

This is a good suggestion: we have now included a Bayesian repeated measures ANOVA to the analyses of Experiment 1. As expected, these analyses provide further, though mild evidence in support for the null hypothesis (See Table S2). For completeness, the new Bayesian analyses are presented alongside the previous frequentist ones in the revised manuscript.

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