

TDAB01 Probability and Statistics

Maryna Prus
IDA, Linköping University

Lecture 6: Stochastic Processes

Overview

- **Definitions and classifications**
- **Markov chains**
- **Binomial processes**
- **Poisson processes**

Stochastic Processes

- ▶ **Stochastic process:**

Sequence of random variables X_1, X_2, \dots **observed over time**

- ▶ Example: X_t = number of "likes" for video on YouTube during day t
- ▶ Example: X_t = temperature in certain city at time t

- ▶ **Stochastic process:**

Random variable $X(t, \omega)$ which also depends on time, where

- ▶ $t \in \mathcal{T}$ and \mathcal{T} - set of times,
can be discrete, e. g. $\mathcal{T} = \{1, 2, 3, \dots\}$ or continuous, e. g. $\mathcal{T} = [0, T]$
- ▶ $\omega \in \Omega$ - outcome of an experiment
- ▶ Values of $X(t, \omega)$ are called **states**
- ▶ Simplified notation $X(t) = X(t, \omega)$

- ▶ Classification of processes:

- ▶ Discrete or continuous **states**
- ▶ Discrete or continuous **time**
- ▶ Discrete states, continuous time: Number of "likes" on YouTube over time
- ▶ Discrete states, discrete time: Number of "likes" on YouTube during day t
- ▶ Continuous states, discrete time: Highest temperature on certain day
- ▶ Continuous states, continuous time: Temperature over time

Markov processes

- ▶ **Markov process:** The forecast for tomorrow depends only on today:

$$P(\text{future}|\text{now, history}) = P(\text{future}|\text{now})$$

- ▶ (Discrete-time) **Markov process:**

For all times $t_1 < \dots < t_n < t_{n+1}$ it holds that

$$P(X(t_{n+1}) = x_{n+1} | X(t_1) = x_1, \dots, X(t_n) = x_n) = P(X(t_{n+1}) = x_{n+1} | X(t_n) = x_n)$$

i.e. $X(t_{n+1})$ is independent of $X(t_1), \dots, X(t_{n-1})$ if $X(t_n)$ is given

- ▶ Well-developed theory, simple techniques
- ▶ But many processes are not Markov

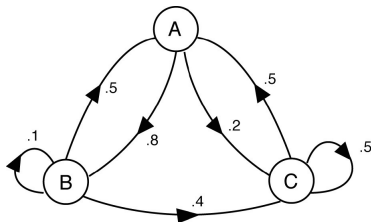
Markov chains

- ▶ **Markov chain**: Markov process with **discrete time** and **discrete states**
- ▶ Time: $\mathcal{T} = \{1, 2, 3, \dots\}$
- ▶ Enumerate states: $1, 2, \dots, n$ (or A, B, C, \dots)
 $n = \infty$ generally possible; not in this course
- ▶ **Transition probability (one-step)**

$$p_{ij}(t) = \mathbf{P}\{X(t+1) = j | X(t) = i\}$$

→ probability to move from state i to state j

- ▶ Example:



Markov chains

- ▶ **Transition probability (h-steps)**

$$p_{ij}^{(h)}(t) = \mathbf{P} \{X(t+h) = j | X(t) = i\}$$

→ probability to move from state i to state j by means of h transitions

- ▶ **Homogeneous Markov chain:** Transition probabilities independent of time:

$$p_{ij}(t) = p_{ij}$$

- ▶ **Transition matrix**

$$P = \begin{pmatrix} p_{11} & p_{12} \\ p_{21} & p_{22} \end{pmatrix}$$

- ▶ Example: $\Omega = \{\text{sunny, rainy}\}$

$P(\text{sunny tomorrow} | \text{sunny today}) = 0.9$, $P(\text{rainy tomorrow} | \text{rainy today}) = 0.3$

- ▶ Example from previous slide

$$P = \begin{pmatrix} 0 & 0.8 & 0.2 \\ 0.5 & 0.1 & 0.4 \\ 0.5 & 0 & 0.5 \end{pmatrix}$$

Markov chains

- ▶ **Transition matrix** 1-step

$$P = \begin{pmatrix} p_{11} & p_{12} & \cdots & p_{1n} \\ p_{21} & p_{22} & \cdots & p_{2n} \\ \vdots & \vdots & \ddots & \vdots \\ p_{n1} & p_{n2} & \cdots & p_{nn} \end{pmatrix}$$

Note that $p_{i1} + p_{i2} + \dots + p_{in} = 1$, for all i

- ▶ **Transition probability (h-steps)**

$$p_{ij}^{(h)}(t) = \mathbf{P}\{X(t+h) = j | X(t) = i\}$$

- ▶ Complicated: many paths for $i \rightarrow j$ when $h > 1$
 - ▶ Example: $\Omega = \{1, 2\}$. If $h = 2$, we can make the trip $1 \rightarrow 2$ in two ways:
 - ▶ $1 \rightarrow 1 \rightarrow 2$
 - ▶ $1 \rightarrow 2 \rightarrow 2$
- Use matrices

- ▶ **Transition matrix** h -steps

$$P^{(h)} = \begin{pmatrix} p_{11}^{(h)} & p_{12}^{(h)} & \cdots & p_{1n}^{(h)} \\ p_{21}^{(h)} & p_{22}^{(h)} & \cdots & p_{2n}^{(h)} \\ \vdots & \vdots & \ddots & \vdots \\ p_{n1}^{(h)} & p_{n2}^{(h)} & \cdots & p_{nn}^{(h)} \end{pmatrix}$$

Note that $p_{i1}^{(h)} + p_{i2}^{(h)} + \dots + p_{in}^{(h)} = 1$, for all i

- ▶ Relation between P and $P^{(h)}$

$$P^{(h)} = P \cdot P \dots P = P^h$$

- ▶ Example: Example 6.9 in textbook

Distribution of $X(h)$

- ▶ **Initial distribution** of $X(t)$ at $t = 0$ is the row vector:

$$P_0 = (P_0(1), P_0(2), \dots, P_0(n)), \quad P_0(i) = P(X(0) = i)$$

- ▶ **Probability distribution over the states** after h steps:

$$P_h = (P_h(1), P_h(2), \dots, P_h(n))$$

- ▶ Computing P_h :

$$P_h = P_0 P^{(h)} = P_0 P^h$$

- ▶ Example: $P_0 = (1/3, 1/3, 1/3)$ and

$$P = \begin{pmatrix} 0.8 & 0.1 & 0.1 \\ 0.1 & 0.6 & 0.3 \\ 0.1 & 0.5 & 0.4 \end{pmatrix}$$

$$\Rightarrow P_3 = (1/3, 1/3, 1/3) \begin{pmatrix} 0.8 & 0.1 & 0.1 \\ 0.1 & 0.6 & 0.3 \\ 0.1 & 0.5 & 0.4 \end{pmatrix}^3 = (0.333, 0.407, 0.259)$$

$\Rightarrow P(X(3) = 1) = 0.333$ - probability for first state after 3 transitions

- ▶ See Example 6.10 in textbook

Steady-state distribution

- ▶ Probability distribution over states after many steps - ?
- ▶ **Steady-state distribution** is the row vector

$$\pi = \lim_{h \rightarrow \infty} P_h$$

- ▶ $\lim_{h \rightarrow \infty} P_h = \lim_{h \rightarrow \infty} P_{h+1} \Rightarrow \pi P = \pi$
- ▶ Computing π :

$$\begin{cases} \pi P = \pi \\ \sum_x \pi_x = 1 \end{cases}$$

- ▶ A Markov chain is **regular** if there is h such that

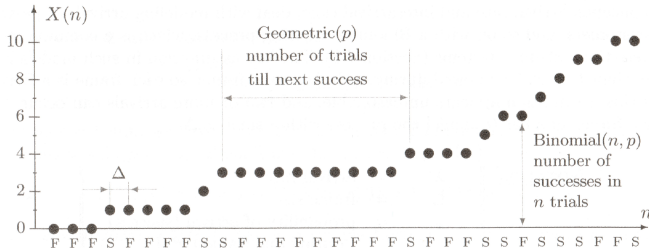
$$p_{ij}^{(h)} > 0$$

for all i, j .

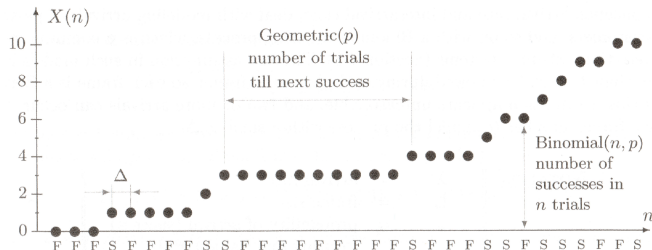
- ▶ **Any regular Markov chain has steady-state distribution**

Binomial process

- ▶ **Counting processes:** $X(t)$ - number of items counted by time t
- ▶ **Binomial process:** $X(n)$ - number of successes in first n independent Bernoulli trials
- ▶ $X(n) \sim \text{Binomial}(n, p)$, p - probability of success
- ▶ Y = number of trials between two consecutive successes
- ▶ $Y \sim \text{Geo}(p)$



Binomial process



- ▶ New Bernoulli trial every Δ seconds $\rightarrow \Delta =$ **time frame**
- ▶ n trials occur during $t = n\Delta$ seconds $\rightarrow n = t/\Delta$
- ▶ Process as function of time: $X(n) = X(t/\Delta)$
- ▶ **Expected number of successes during t seconds:**

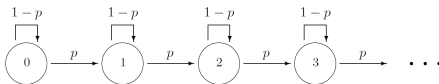
$$\mathbb{E}(X(t/\Delta)) = tp/\Delta$$

- ▶ Expected number of successes per second: $\lambda = p/\Delta$
- ▶ λ is also called **arrival rate**

Binomial process

- ▶ **Interarrival time** T - **time** between two consecutive successes
- ▶ Y = **number** of trials between two consecutive successes, $Y \sim \text{Geo}(p)$
- ▶ Interarrival time: $T = Y\Delta$
- ▶ T has rescaled geometric distribution with support $\Delta, 2\Delta, 3\Delta, \dots$
→ $\mathbb{E}(T) = \mathbb{E}(Y\Delta) = \Delta\mathbb{E}(Y) = \Delta/p = 1/\lambda$
- ▶ Binomial process - homogeneous Markov chain with transition probabilities

$$p_{ij} = \begin{cases} p, j = i + 1 \\ 1 - p, j = i \\ 0, \text{otherwise} \end{cases}$$



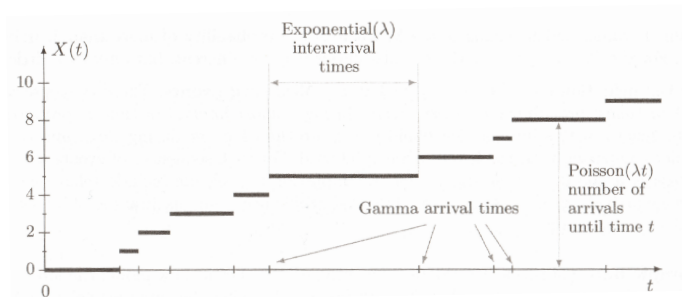
This Markov chain is non-regular as $X(n)$ non-decreasing, $p_{i,i-1}^{(h)} = 0$ for all h

→ no steady-state distribution

- ▶ Example: Example 6.18 in textbook

Poisson process

- ▶ **Poisson process** - continuous-time stochastic process obtained from Binomial process by letting $\Delta \rightarrow 0$ while keeping λ constant
- ▶ From Lecture 3: $X(t) \sim \text{Binomial}(t/\Delta, p) \rightarrow \text{Poisson}(\lambda t)$ when $n = t/\Delta \rightarrow \infty$ and $p = \lambda\Delta \rightarrow 0$



- ▶ Interarrival time $T \sim \text{Exp}(\lambda)$
- ▶ Time for k -th success: $T_k \sim \text{Gamma}(k, \lambda)$ - sum of k interarrival times
- ▶ Example: See Example 6.20 in textbook

Simulation of stochastic processes

- ▶ For simulation of stochastic processes
see Chapter 6.4 in textbook

- ▶ See also

`SimulateMarkovChain.R`

for simulation Markov chains

`SimulateBinomialProcess.R`

for simulation of Binomial processes

`SimulatePoissonProcess.R`

for simulation of Poisson processes

Thank you for your attention!