Efficient Likelihood-Free Inference

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Likelihood-free inference:

Perform statistical inference for models where

- 1. the likelihood function is too costly to evaluate
- 2. sampling simulating data from the model is possible

Importance

Such models and inference problems occur widely

- Neuroscience: Simulating neural circuits
- Evolutionary biology: Simulating evolution

- Computer vision: Simulating naturalistic scenes
- Health science: Simulating the spread of an infectious disease



Background

Previous work

Our approach

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Assumptions on the models

- Only assumption: sampling simulating data from the model is possible
- Models specified by a data generating mechanism
 - e.g. stochastic nonlinear dynamical systems
 - e.g. computer models / simulators of some complex physical or biological process
- Different communities use different names:
 - Simulator-based models
 - Stochastic simulation models
 - Implicit models
 - Generative (latent-variable) models
 - Probabilistic programs

Definition of simulator-based models (SBMs)

- Let $(\Omega, \mathcal{F}, \mathcal{P})$ be a probability space.
- A simulator-based model is a collection of (measurable) functions g(., θ) parametrised by θ,

$$\boldsymbol{\omega} \in \Omega \mapsto \boldsymbol{x}_{\boldsymbol{\theta}} = \boldsymbol{g}(\boldsymbol{\omega}, \boldsymbol{\theta}) \in \mathcal{X} \tag{1}$$

- For any fixed θ , $\mathbf{x}_{\theta} = g(., \theta)$ is a random variable.
- $g(., \theta)$ typically not available in closed form



- Direct implementation of hypotheses of how the observed data were generated.
- Neat interface with scientific models (e.g. from physics or biology).
- Modelling by replicating the mechanisms of nature that produced the observed/measured data. ("Analysis by synthesis")
- Possibility to perform experiments in silico.

- Generally elude analytical treatment.
- Can be easily made more complicated than necessary.
- Statistical inference is difficult.

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Main reason: Likelihood function is too expensive to evaluate

The likelihood function $L(\theta)$

- Well defined but generally intractable for SBMs
- Probability that the model generates data like x^o when using parameter value θ



Three foundational issues

- 1. How should we assess whether $x_{\theta} \equiv x^{o}$?
- 2. How should we compute the probability of the event $x_{\theta} \equiv x^{o}$?
- 3. For which values of θ should we compute it?



Likelihood: Probability that the model generates data like x^o for parameter value heta

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Approximate Bayesian computation

For recent review, see: Lintusaari et al (2017) "Fundamentals and recent developments in approximate Bayesian computation", Systematic Biology

- 1. How should we assess whether $\mathbf{x}_{\theta} \equiv \mathbf{x}^{o}$?
 - \Rightarrow Check whether $||T(\mathbf{x}_{\theta}) T(\mathbf{x}^{o})|| \leq \epsilon$
- 2. How should we compute the proba of the event $x_{\theta} \equiv x^{o}$? \Rightarrow By counting
- 3. For which values of θ should we compute it?
 - \Rightarrow Sample from the prior (or other proposal distributions)

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 \Rightarrow Sample from the prior (or other proposal distributions) Difficulties:

- Choice of T() and e
- Typically high computational cost

Implicit likelihood approximation





$$\mathcal{L}(oldsymbol{ heta}) pprox \mathbb{P}^{N}(d(oldsymbol{x}_{oldsymbol{ heta}},oldsymbol{x}^{o}) \leq \epsilon) = rac{1}{N}\sum_{i=1}^{N}\mathbb{1}\left(d(oldsymbol{x}_{oldsymbol{ heta}}^{(i)},oldsymbol{x}^{o}) \leq \epsilon
ight)$$

Synthetic likelihood

(Simon Wood, Nature, 2010)

- 1. How should we assess whether $\mathbf{x}_{\theta} \equiv \mathbf{x}^{o}$?
- 2. How should we compute the proba of the event $x_{\theta} \equiv x^{o}$?
 - \Rightarrow Compute summary statistics $\boldsymbol{t}_{\boldsymbol{\theta}} = T(\boldsymbol{x}_{\boldsymbol{\theta}})$
 - \Rightarrow Model their distribution as a Gaussian
 - \Rightarrow Compute likelihood function with $T(\mathbf{x}^o)$ as observed data
- 3. For which values of θ should we compute it?
 - ⇒ Use obtained "synthetic" likelihood function as part of a Monte Carlo method

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Difficulties:

- ► Choice of *T*()
- Gaussianity assumption may not hold
- Typically high computational cost

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1. How should we assess whether $\mathbf{x}_{\theta} \equiv \mathbf{x}^{o}$?

 \Rightarrow Use classification (Gutmann et al, 2014, 2017)

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- 2. How should we compute the proba of the event $x_{\theta} \equiv x^{o}$?

 \Rightarrow Use density ratio estimation (Dutta et al, 2016, arXiv:1611.10242)

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 \Rightarrow Use Bayesian optimisation (Gutmann and Corander, 2013-2016)

Overview of some of my work

- 1. How should we assess whether $x_{\theta} \equiv x^{o}$?
 - \Rightarrow Use classification (Gutmann et al, 2014, 2017)
- Basic idea: Classification accuracy (discriminability) serves as distance measure
- Value of 1: far; Value of 1/2: close



Overview of some of my work

- 1. How should we assess whether $x_{\theta} \equiv x^{o}$?
- 2. How should we compute the proba of the event $x_{\theta} \equiv x^{\circ}$?
 - \Rightarrow Use density ratio estimation (Dutta et al, 2016, arXiv:1611.10242)
- Basic idea: frame posterior estimation as ratio estimation problem

$$p(\theta|\mathbf{x}) = \frac{p(\mathbf{x}|\theta)p(\theta)}{p(\mathbf{x})} = r(\mathbf{x},\theta)p(\theta)$$
(2)

Estimate r̂(x, θ) yields estimate of the likelihood function and posterior

$$\hat{L}(\boldsymbol{\theta}) \propto \hat{r}(\boldsymbol{x}^{o}, \boldsymbol{\theta}), \qquad \hat{p}(\boldsymbol{\theta}|\boldsymbol{x}^{o}) = \hat{r}(\boldsymbol{x}^{o}, \boldsymbol{\theta})p(\boldsymbol{\theta}).$$
 (3)

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Why is the ABC algorithm so expensive?

- 1. It rejects most samples when ϵ is small
- 2. It does not make assumptions about the shape of $L(\theta)$
- 3. It does not use all information available
- 4. It aims at equal accuracy for all parameters



Proposed solution

(Gutmann and Corander, JMLR, 2016)

- 1. It rejects most samples when ϵ is small \Rightarrow Don't reject samples – learn from them
- 2. It does not make assumptions about the shape of $L(\theta)$ \Rightarrow Model the distances, assume average distance is smooth
- It does not use all information available
 ⇒ Use Bayes' theorem to update the model
- 4. It aims at equal accuracy for all parameters
 ⇒ Prioritize parameter regions with small distances

equivalent strategy applies to inference with synthetic likelihood

Modelling (points 1 & 2)

- Data are tuples (θ_i, d_i) , where $d_i = d(\mathbf{x}_{\theta}^{(i)}, \mathbf{x}^o)$
- Model the conditional distribution of d given heta
- Estimated model yields approximation $\hat{L}(\boldsymbol{\theta})$ for any choice of ϵ

$$\hat{L}(\boldsymbol{ heta}) \propto \widehat{\mathbb{P}} \left(\boldsymbol{d} \leq \epsilon \mid \boldsymbol{ heta}
ight)$$

 $\widehat{\mathbb{P}}$ is probability under the estimated model.

- Here: Use (log) Gaussian process with squared exponential covariance function as model
- Approach not restricted to this model or Gaussian processes (comparison of different GP models: Järvenpää et al, 2016, arXiv:1610.06462)

Data acquisition (points 3 & 4)

- Samples of θ could be obtained by sampling from the prior or some adaptively constructed proposal distribution
- Give priority to regions in the parameter space where distance d tends to be small.
- Use Bayesian optimization to find such regions
- ► Here: Use lower confidence bound acquisition function (e.g. Cox and John, 1992; Srinivas et al, 2012)

$$\mathcal{A}_{t}(\theta) = \underbrace{\mu_{t}(\theta)}_{\text{post mean}} - \sqrt{\underbrace{\eta_{t}^{2}}_{\text{weight post var}}} \underbrace{v_{t}(\theta)}_{\text{weight post var}}$$
(4)

t: number of samples acquired so far

 Approach not restricted to this acquisition function. (new acquisition function: Järvenpää et al, 2017, arXiv:1704.00520)

Bayesian optimization for likelihood-free inference



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Example: Bacterial infections in child care centers

- Likelihood intractable for cross-sectional data
- But generating data from the model is possible



Example: Bacterial infections in child care centers

- Comparison of the proposed approach with a standard population Monte Carlo ABC approach.
- Roughly equal results using 1000 times fewer simulations.



(Gutmann and Corander, JMLR, 2016)

Example: Bacterial infections in child care centers

- Comparison of the proposed approach with a standard population Monte Carlo ABC approach.
- ▶ Roughly equal results using 1000 times fewer simulations.



Posterior means are shown as solid lines, credibility intervals as shaded areas or dashed lines.

Benefits

- The proposed method makes the inference more efficient.
 - allowed us to perform far more comprehensive data analysis than with standard approach (Numminen et al, 2016)
- Enables inference for models which were out of reach till now
 - model of evolution where simulating a single data set took us 12-24 hours (Marttinen et al, 2015)
- Enables easier assessment of parameter identifiability for complex models
 - model about transmission dynamics of tuberculosis (Lintusaari et al, 2016)

- Model: How to best model the distance between simulated and observed data?
- Acquisition function: Can we find strategies which are optimal for parameter inference?
- Efficient high-dimensional inference: Can we use the approach to infer the joint distribution of 1000 variables?

see Gutmann and Corander, JMLR, 2016 for a discussion

for first answers: http://homepages.inf.ed.ac.uk/mgutmann

- Topic: Inference for models where the likelihood is intractable but sampling is possible
- Inference principle: Find parameter values for which the distance between simulated and observed data is small
- Problem considered: Computational cost
- Proposed approach: Combine statistical modeling of the distance with decision making under uncertainty (Bayesian optimization)
- Outcome: Approach increases the efficiency of the inference by several orders of magnitude